

# Convex ordering for insurance preferences

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## Problem Formulation

Let  $X$  be a non-negative random variable to model an insurable loss of an agent. After purchasing an insurance  $I$ , the policyholder retains a part of the loss  $R(X) \triangleq X - I(X)$ , and pays a premium  $P(I(X))$ . The total loss exposure of the policyholder is

$$L_I(X) \triangleq X - I(X) + P(I(X)) = R(X) + P(I(X)).$$

## Problem Formulation (cont.)

Consider the admissible set  $\mathcal{I}$  of ceded loss functions  $I$ :

$$\mathcal{I} \triangleq \{I : [0, \text{ess sup } X] \rightarrow [0, \text{ess sup } X]\}$$

$$(I1) : 0 \leq I(x) \leq x, \quad \text{for any } x \in [0, \text{ess sup } X];$$

$$(I2) : 0 \leq I(x_1) - I(x_2) \leq x_1 - x_2, \quad \text{for any } 0 \leq x_2 < x_1 \leq \text{ess sup } X\}.$$

Also, consider the admissible set  $\mathcal{R}$  of retained loss functions  $R \triangleq Id - I$ :

$$\mathcal{R} \triangleq \{R : [0, \text{ess sup } X] \rightarrow [0, \text{ess sup } X]\}$$

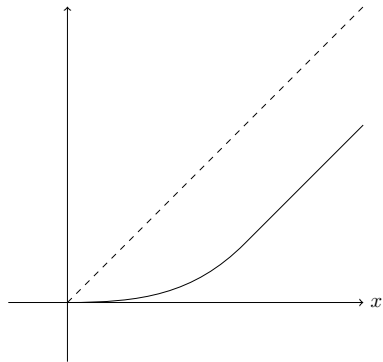
$$(R1) : 0 \leq R(x) \leq x, \quad \text{for any } x \in [0, \text{ess sup } X];$$

$$(R2) : 0 \leq R(x_1) - R(x_2) \leq x_1 - x_2, \quad \text{for any } 0 \leq x_2 < x_1 \leq \text{ess sup } X\}.$$

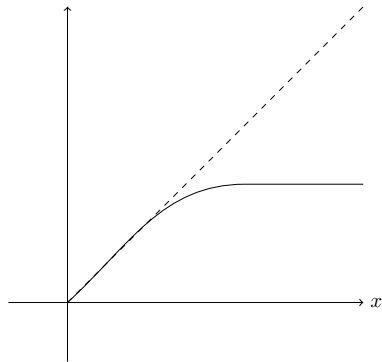
Both admissible  $I$  and  $R$  are non-decreasing.

# Admissible $I$ and $R$

$$I(x) = x - R(x)$$

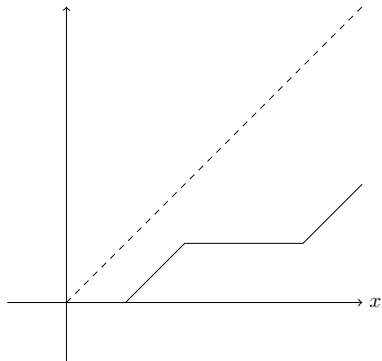


$$R(x) = x - I(x)$$

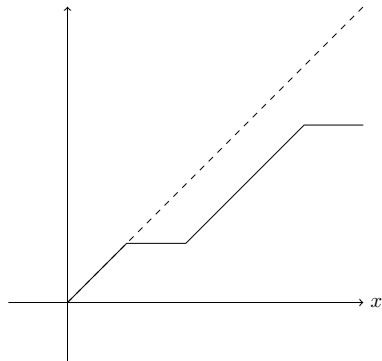


# Admissible $I$ and $R$

$$I(x) = x - R(x)$$

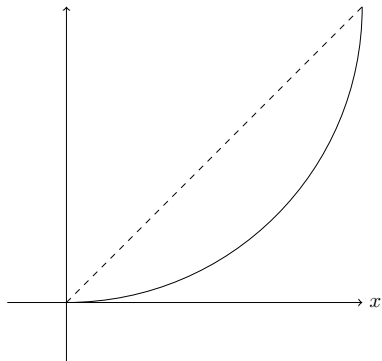


$$R(x) = x - I(x)$$

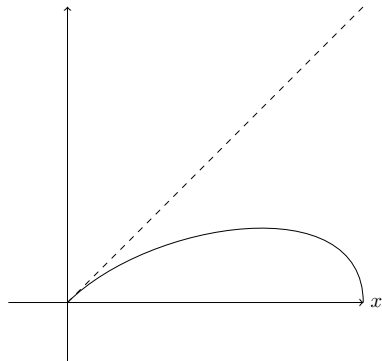


# Non-admissible $I$ and $R$

$$I(x) = x - R(x)$$

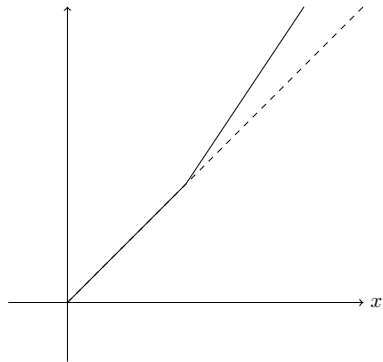


$$R(x) = x - I(x)$$

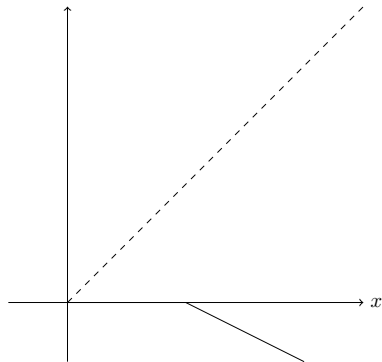


# Non-admissible $I$ and $R$

$$I(x) = x - R(x)$$



$$R(x) = x - I(x)$$



## Problem Formulation (cont.)

**Problem:** Let  $G$  be a non-decreasing (in all arguments) function from  $\mathbb{R}_+^3$  to  $\mathbb{R}$ . Let  $g$  be a non-decreasing (in both arguments) function from  $\mathbb{R}_+^2$  to  $\mathbb{R}_+$ . Let  $V_{\text{cx}}$  be a functional preserving convex order. Let  $\alpha \in (0, 1]$ . The policyholder aims to choose an optimal insurance indemnity  $I \in \mathcal{I}$  that minimizes

$$G(\mathbb{E}[L_I(X)], \text{V@R}_\alpha(L_I(X)), \text{TV@R}_\alpha(L_I(X))),$$

with

$$P(I(X)) = g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X))).$$

The premium principle includes the expected value principle, Wang's principle, Dutch principle, variance principle, and standard deviation principle, etc.



# Preliminaries

- The (left-continuous) quantile function of  $Y$  is defined by

$$F_Y^{-1}(p) \triangleq \inf\{y \in \mathbb{R} | F_Y(y) \geq p\}, \quad \forall p \in (0, 1].$$

- The Value-at-Risk of  $Y$  at the level  $\alpha \in (0, 1]$ , denoted by  $\text{V@R}_\alpha(Y)$ , is defined by

$$\text{V@R}_\alpha(Y) \triangleq F_Y^{-1}(\alpha).$$

- The Tail Value-at-Risk of  $Y$  at the level  $\alpha \in (0, 1]$ , denoted by  $\text{TV@R}_\alpha(Y)$ , is defined by

$$\text{TV@R}_\alpha(Y) \triangleq \frac{1}{1-\alpha} \int_\alpha^1 \text{V@R}_\lambda(Y) d\lambda = \frac{1}{1-\alpha} \int_\alpha^1 F_Y^{-1}(\lambda) d\lambda,$$

for  $\alpha < 1$ ; when  $\alpha = 1$ ,  $\text{TV@R}_\alpha(Y)$  is defined by the  $\text{ess sup } Y$ .

- The stop-loss transform of  $Y$  is defined by

$$\pi_Y(t) \triangleq \mathbb{E}[(Y - t)_+] = \int_t^{\text{ess sup } Y} S_Y(y) dy, \quad \forall t \in \mathbb{R}.$$

- $Y$  and  $Z$  are comonotonic if there exist a random variable  $S$  and two non-decreasing functions  $f_1$  and  $f_2$  such that  $Y \stackrel{d}{=} f_1(S)$  and  $Z \stackrel{d}{=} f_2(S)$ .

## Proposition 1

Assume that  $\alpha \in (0, 1]$ .

- (i) Both  $V\textcircled{R}$  and  $TV\textcircled{R}$  are translation invariant and comonotonic additive, i.e., for comonotonic  $Y$  and  $Z$ , and  $c \in \mathbb{R}$ ,

$$(T)V\textcircled{R}_\alpha(Y + c) = (T)V\textcircled{R}_\alpha(Y) + c;$$

$$(T)V\textcircled{R}_\alpha(Y + Z) = (T)V\textcircled{R}_\alpha(Y) + (T)V\textcircled{R}_\alpha(Z).$$

- (ii) For any non-decreasing continuous function  $f$ ,

$$V\textcircled{R}_\alpha(f(Y)) = f(V\textcircled{R}_\alpha(Y)).$$

- (iii) For  $\alpha \in (0, 1)$ ,

$$TV\textcircled{R}_\alpha(Y) = V\textcircled{R}_\alpha(Y) + \frac{1}{1 - \alpha} \pi_Y(V\textcircled{R}_\alpha(Y));$$

for  $\alpha = 1$ ,

$$TV\textcircled{R}_\alpha(Y) = V\textcircled{R}_\alpha(Y) = \text{ess sup } X.$$

## Preliminaries (cont.)

By (i) in Proposition 1,

$$\mathbb{E}[L_I(X)] = \mathbb{E}[X - I(X) + P(I(X))] = \mathbb{E}[X] - \mathbb{E}[I(X)] + P(I(X)).$$

$$\mathbf{V}\textcircled{\mathbf{R}}_{\alpha}(L_I(X)) = \mathbf{V}\textcircled{\mathbf{R}}_{\alpha}(R(X) + P(I(X))) = \mathbf{V}\textcircled{\mathbf{R}}_{\alpha}(X) - \mathbf{V}\textcircled{\mathbf{R}}_{\alpha}(I(X)) + P(I(X)).$$

$$\mathbf{TV}\textcircled{\mathbf{R}}_{\alpha}(L_I(X)) = \mathbf{TV}\textcircled{\mathbf{R}}_{\alpha}(R(X) + P(I(X))) = \mathbf{TV}\textcircled{\mathbf{R}}_{\alpha}(X) - \mathbf{TV}\textcircled{\mathbf{R}}_{\alpha}(I(X)) + P(I(X)).$$

Hence, the problem is equivalent to

$$\begin{aligned} \inf_{I \in \mathcal{I}} G\{ & \mathbb{E}[X] - \mathbb{E}[I(X)] + g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X))), \\ & \mathbf{V}\textcircled{\mathbf{R}}_{\alpha}(X) - \mathbf{V}\textcircled{\mathbf{R}}_{\alpha}(I(X)) + g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X))), \\ & \mathbf{TV}\textcircled{\mathbf{R}}_{\alpha}(X) - \mathbf{TV}\textcircled{\mathbf{R}}_{\alpha}(I(X)) + g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X)))\}. \end{aligned}$$

## Preliminaries (cont.)

- $Y$  is smaller than  $Z$  in convex order sense if, for all convex functions  $\varphi : \mathbb{R} \rightarrow \mathbb{R}$ ,  $\mathbb{E}[\varphi(Y)] \leq \mathbb{E}[\varphi(Z)]$ . The convex ordering is denoted by  $Y \leq_{\text{cx}} Z$ .
- $V$  preserves convex order if  $V(Y) \leq V(Z)$  for  $Y \leq_{\text{cx}} Z$ .
- The distribution functions  $F_Y$  and  $F_Z$  cross  $k \geq 1$  times if there exist

$$\xi_{0,2} < \xi_{1,1} \leq \xi_{1,2} < \xi_{2,1} \leq \xi_{2,2} < \cdots < \xi_{k,1} \leq \xi_{k,2} < \xi_{k+1,1},$$

where

$$\xi_{0,2} \triangleq \inf\{x : F_Y(x) \neq F_Z(x)\},$$

and

$$\xi_{k+1,1} \triangleq \sup\{x : F_Y(x) \neq F_Z(x)\},$$

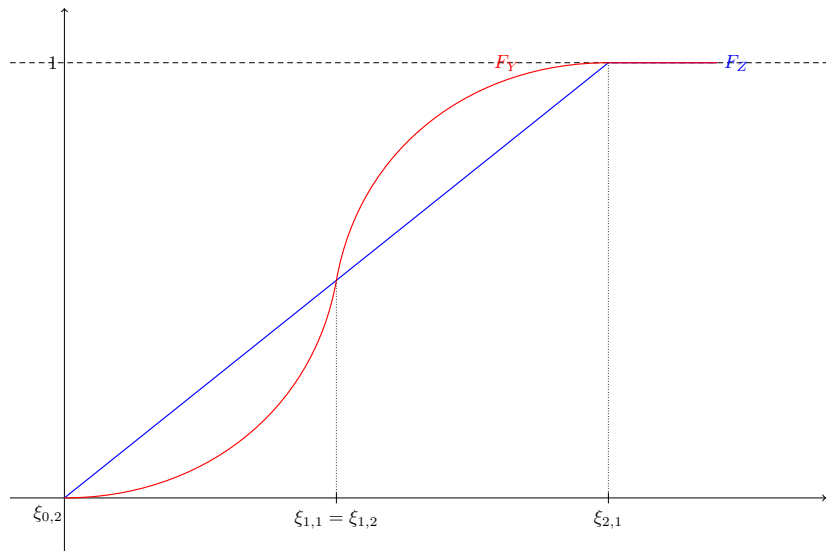
such that, for each  $i = 1, 2, \dots, k$ ,

- (i) for any  $y \in (\xi_{i-1,2}, \xi_{i,1})$  and  $z \in (\xi_{i,2}, \xi_{i+1,1})$ ,

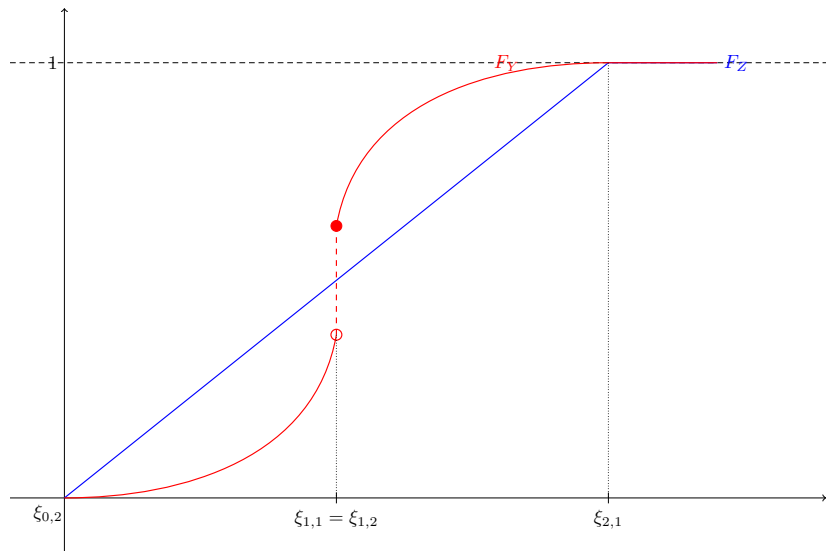
$$(F_Y(y) - F_Z(y))(F_Y(z) - F_Z(z)) < 0; \text{ and}$$

- (ii) if  $\xi_{i,1} < \xi_{i,2}$ ,  $F_Y(z) = F_Z(z)$  for any  $\xi_{i,1} \leq z < \xi_{i,2}$ .

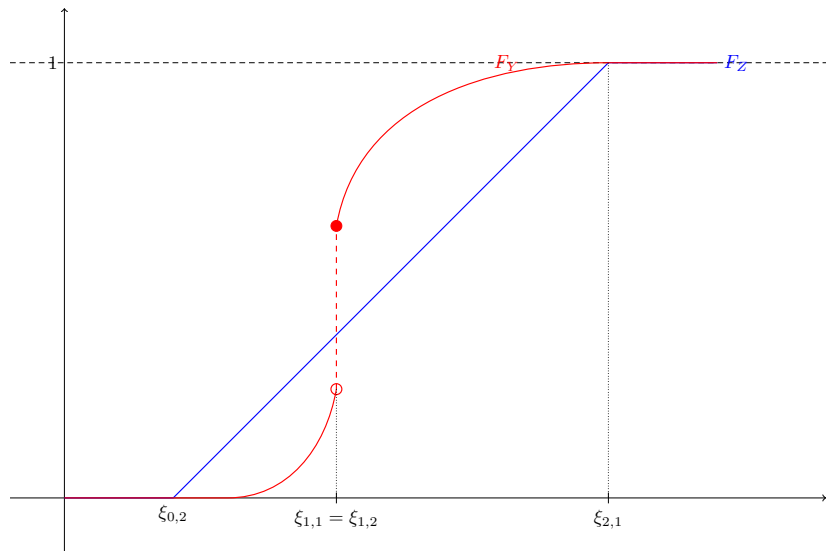
# Crossing Once



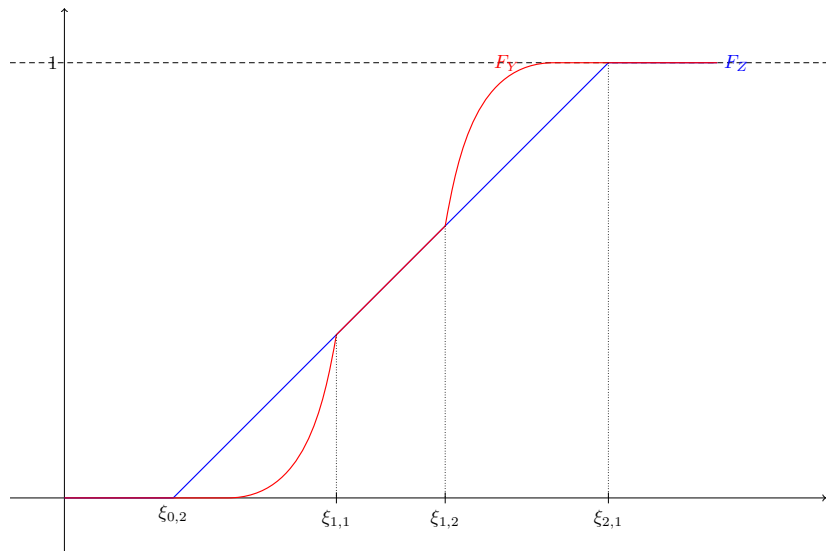
# Crossing Once with Jump



# Crossing Once with Jump

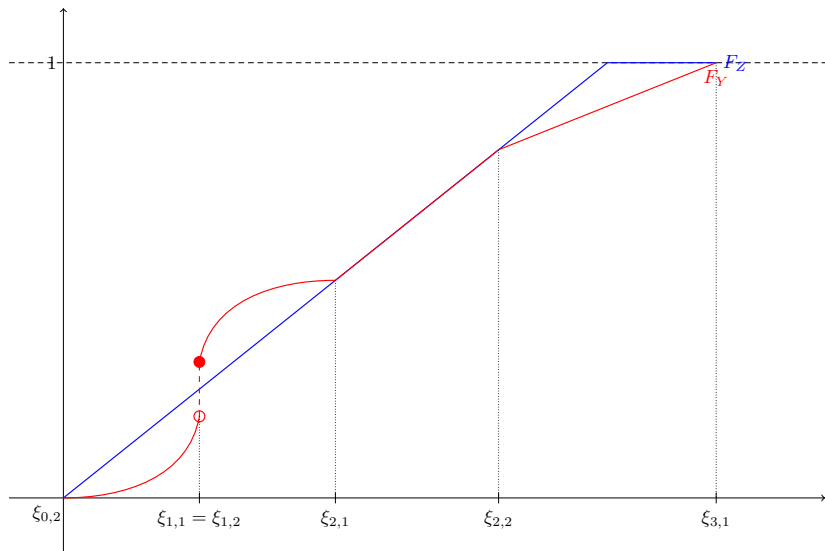


# Crossing Once with Overlapping

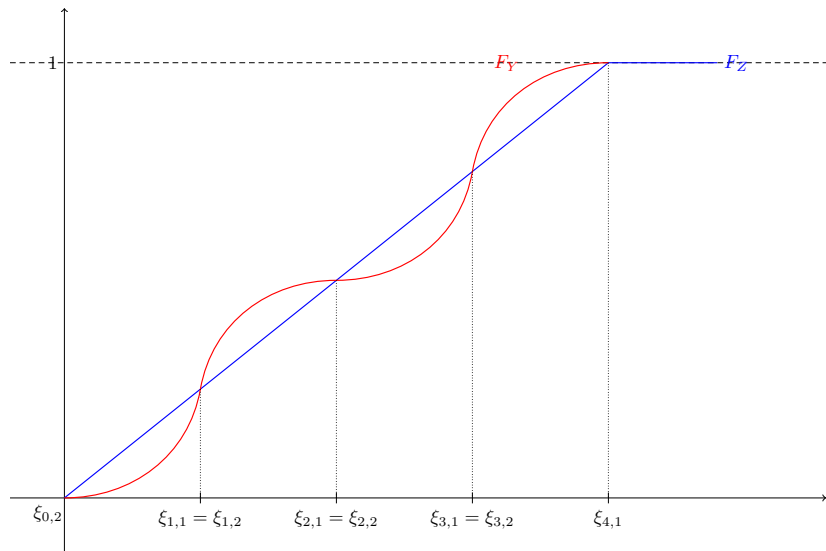




# Crossing Twice with Jump and Overlapping



# Crossing Thrice



## Preliminaries (cont.)

### Theorem 2 (Karlin-Novikoff once-crossing conditions)

Assume that  $F_Y$  and  $F_Z$  cross once. Then  $Y \leq_{cx} Z$  if, and only if,

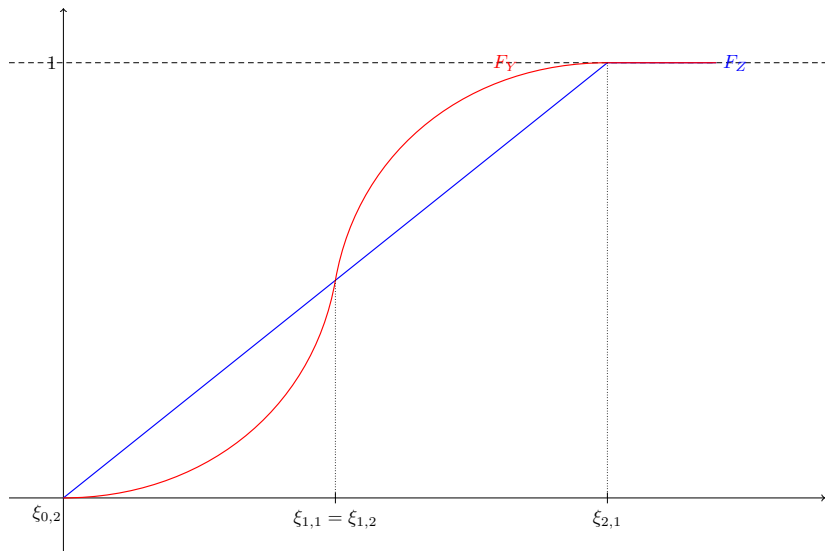
- (i)  $\mathbb{E}[Y] = \mathbb{E}[Z]$ ; and
- (ii)  $F_Y(x) < F_Z(x)$  for  $\xi_{0,2} < x < \xi_{1,1}$ .

### Theorem 3 (Karlin-Novikoff-Stoyan-Taylor crossing conditions)

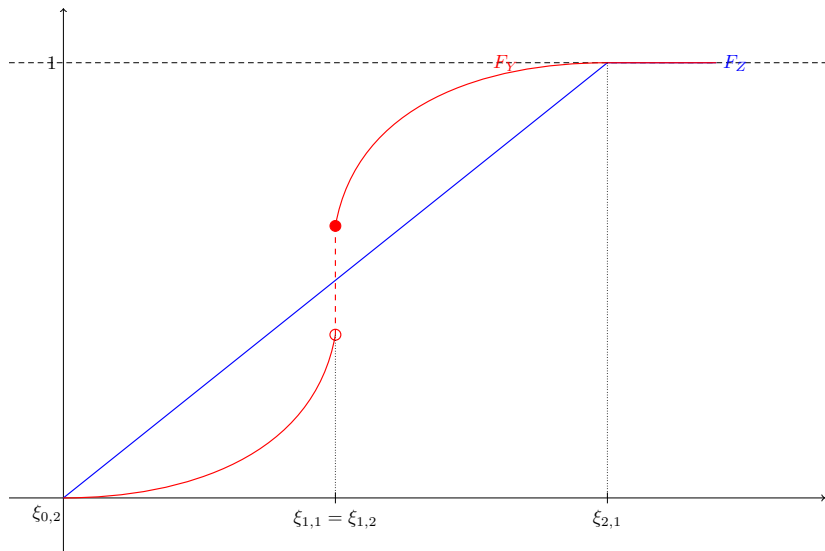
Assume that  $F_Y$  and  $F_Z$  cross  $n \geq 1$  times. Then  $Y \leq_{cx} Z$  if, and only if,

- (i)  $\mathbb{E}[Y] = \mathbb{E}[Z]$ ;
- (ii)  $F_Y(x) < F_Z(x)$  for  $\xi_{0,2} < x < \xi_{1,1}$ ;
- (iii) there is an odd number of crossings  $n = 2m - 1$  where  $m = 1, 2, \dots$ ; and
- (iv) if  $m \geq 2$ , for any  $j = 1, 2, \dots, m - 1$ ,  $\pi_Y(\xi_{2j,2}) \leq \pi_Z(\xi_{2j,2})$ .

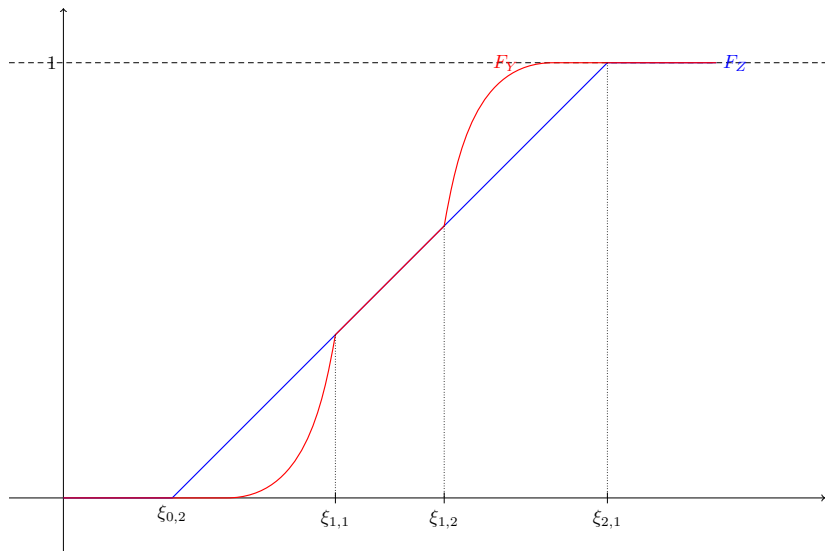
$$Y \leq_{\text{cx}} Z$$



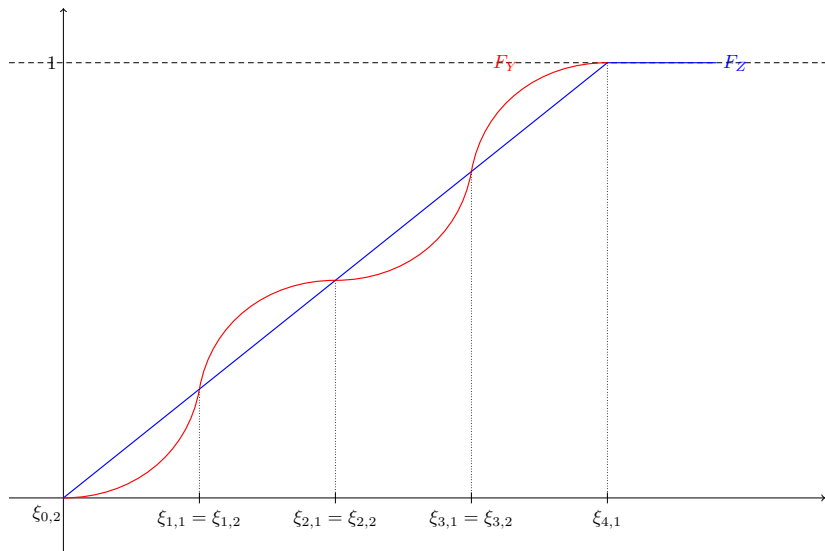
$$Y \leq_{\text{cx}} Z$$



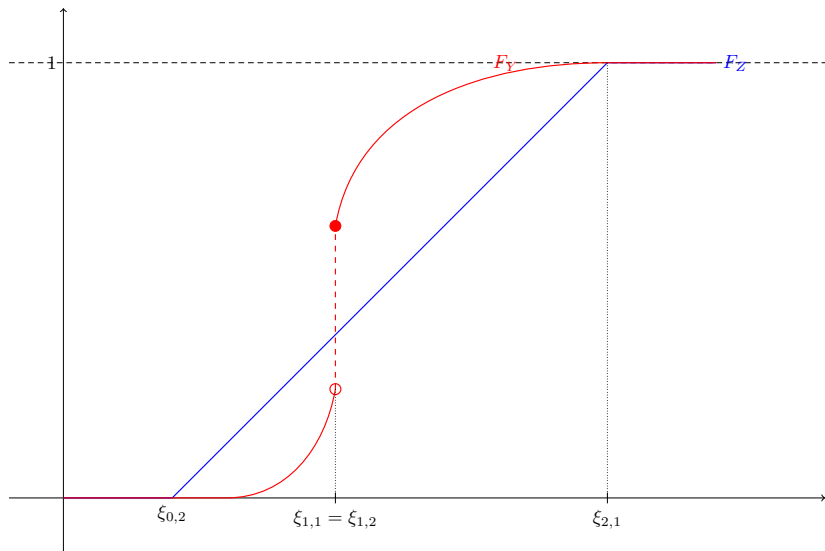
$$Y \leq_{cx} Z$$



$$Y \leq_{\text{cx}} Z$$

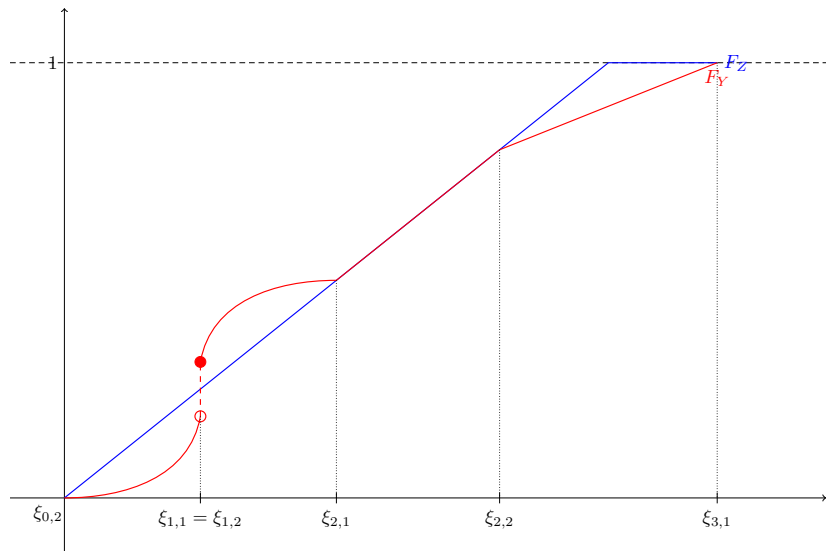


Not  $Y \leq_{cx} Z$





Not  $Y \leq_{cx} Z$



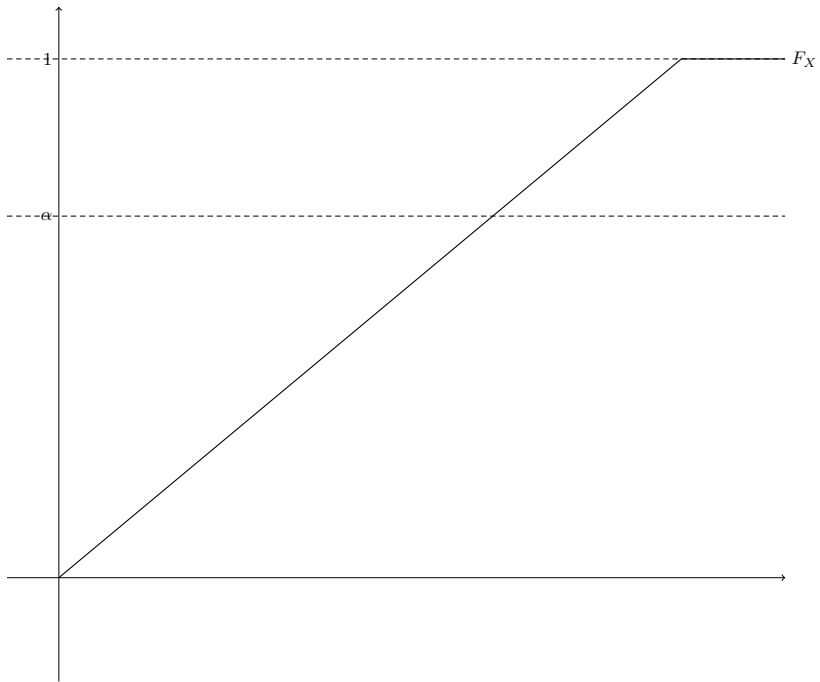
# Distribution Function of $I(X)$ for Admissible $I$

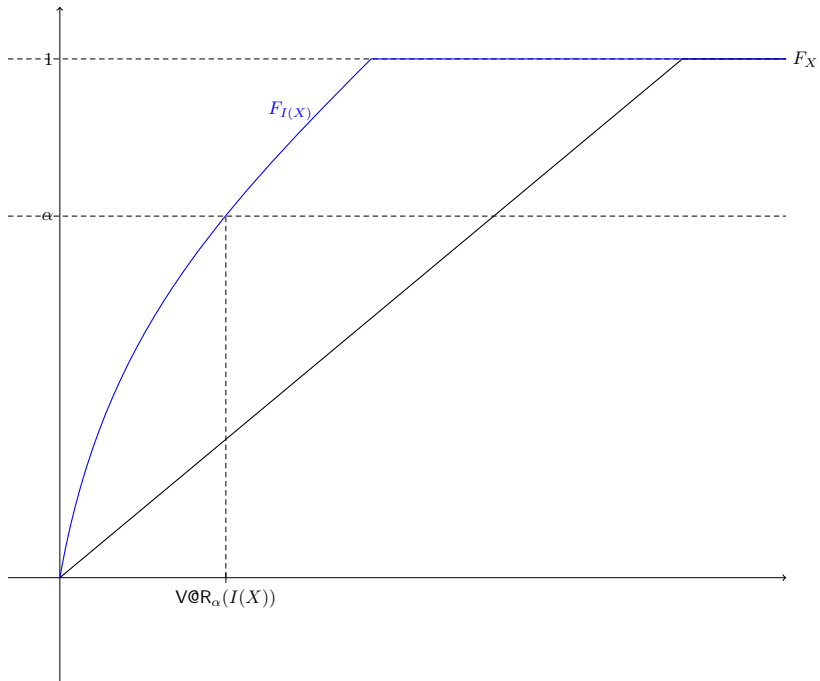
Conditions (I1) and (I2) in  $\mathcal{I}$  can be expressed in terms of the distribution function of  $I(X)$ :

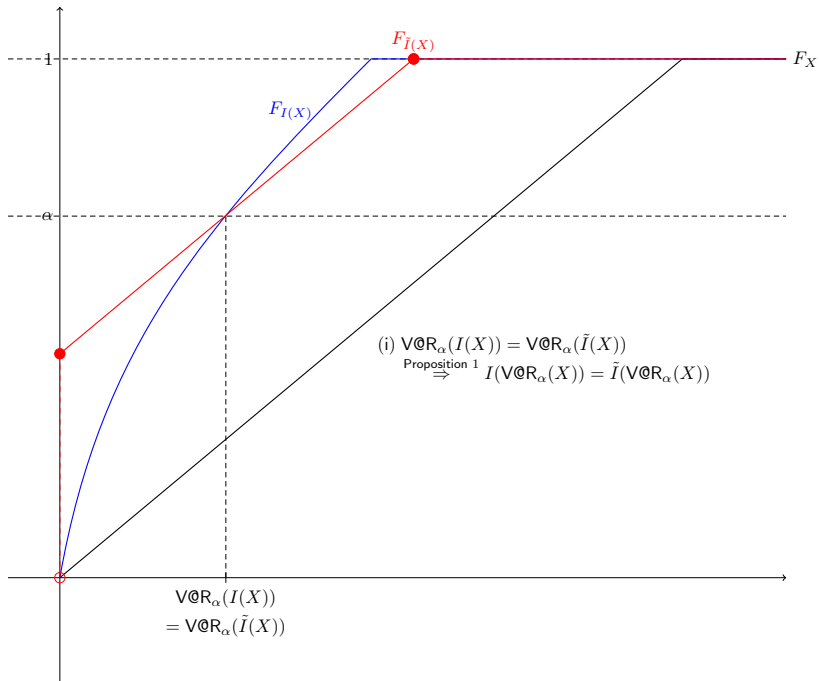
- (i)  $F_{I(X)}(0-) = 0$  and  $F_X(z) \leq F_{I(X)}(z)$  for any  $z \in [0, \text{ess sup } X]$ ;
- (ii)  $0 \leq F_{I(X)}^{-1}(\beta_2) - F_{I(X)}^{-1}(\beta_1) \leq F_X^{-1}(\beta_2) - F_X^{-1}(\beta_1)$  for any  $0 \leq \beta_1 < \beta_2 \leq 1$ .

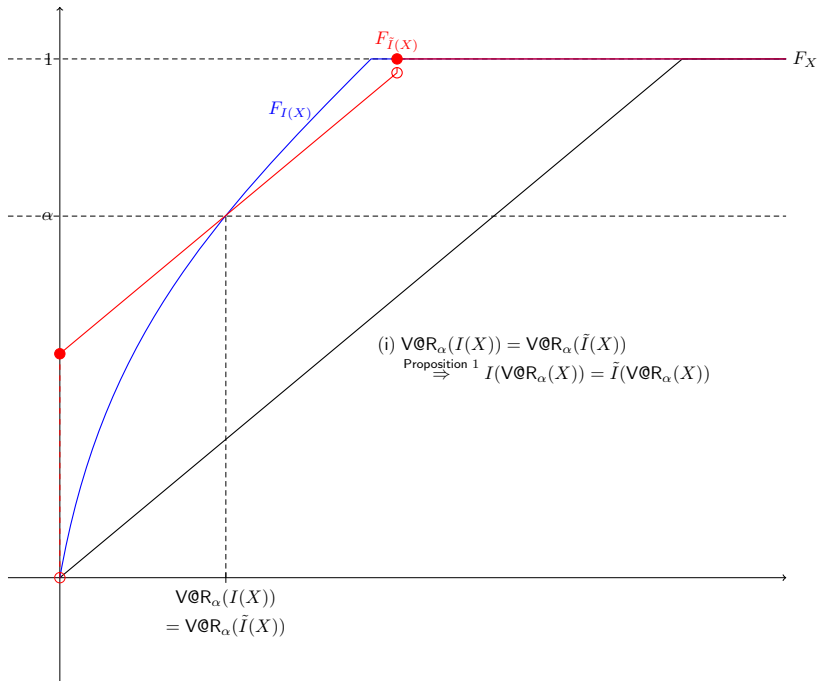
Recall that the problem is

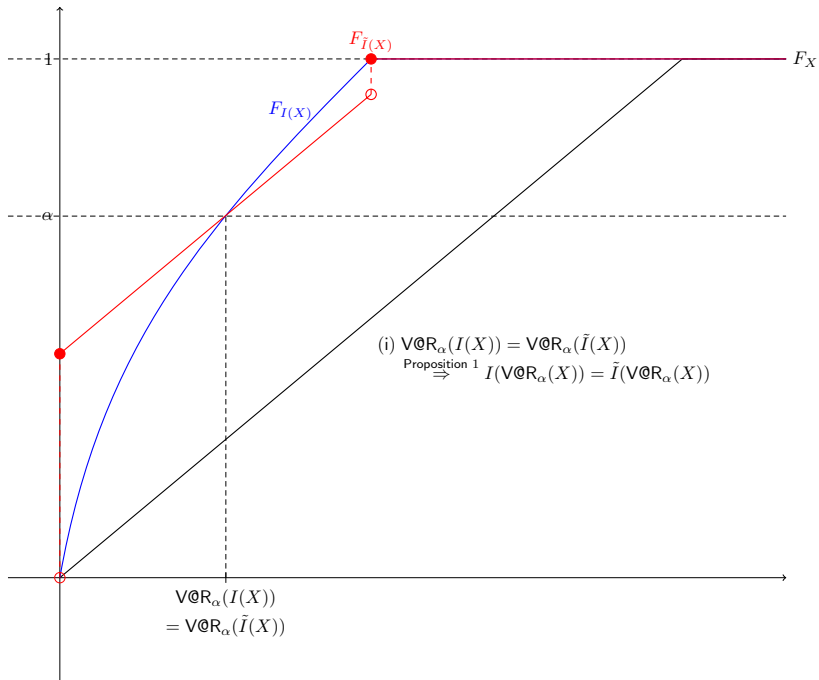
$$\inf_{I \in \mathcal{I}} G\{\mathbb{E}[X] - \mathbb{E}[I(X)] + g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X))), \\ V\text{@}R_{\alpha}(X) - V\text{@}R_{\alpha}(I(X)) + g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X))), \\ \text{TV@}R_{\alpha}(X) - \text{TV@}R_{\alpha}(I(X)) + g(\mathbb{E}[I(X)], V_{\text{cx}}(I(X)))\}.$$

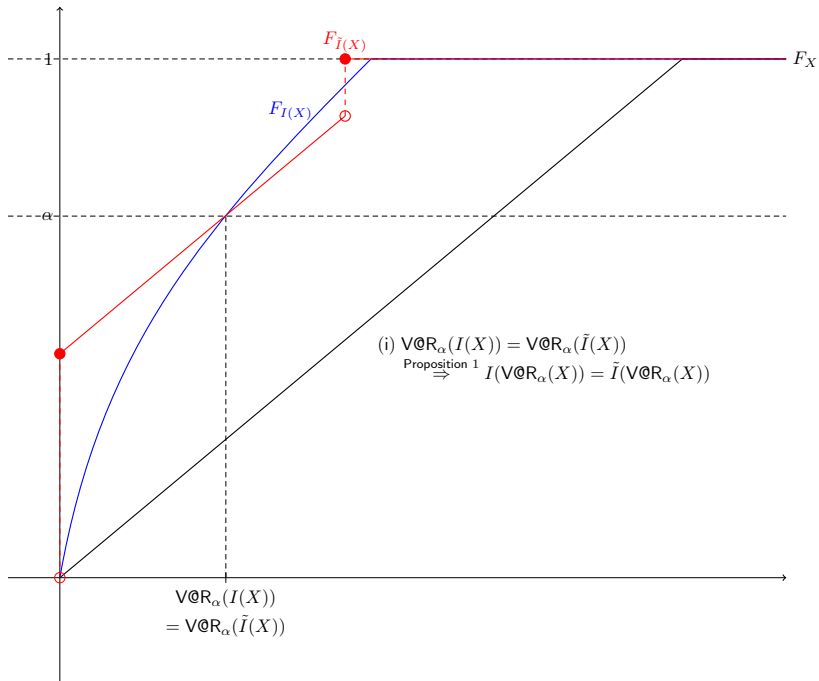




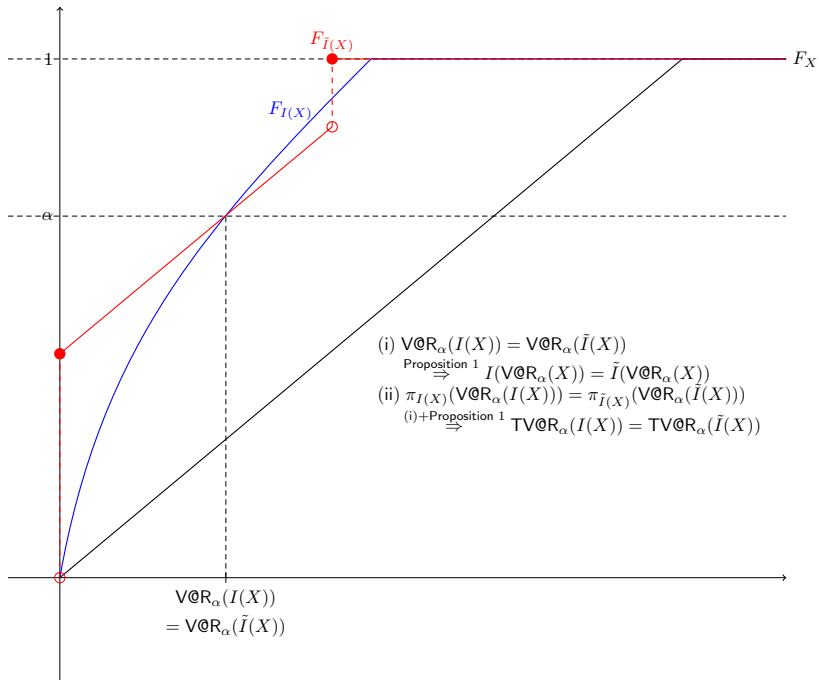


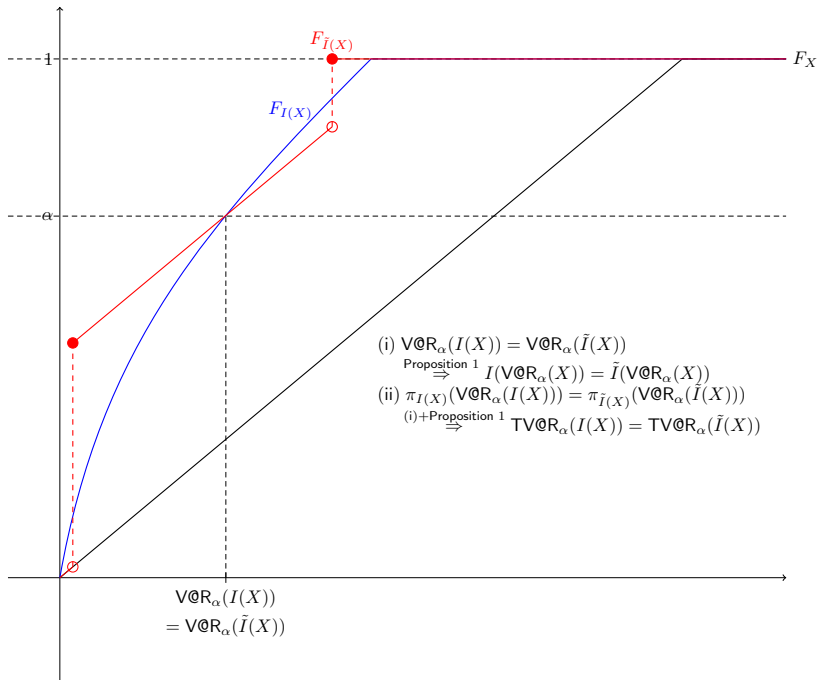


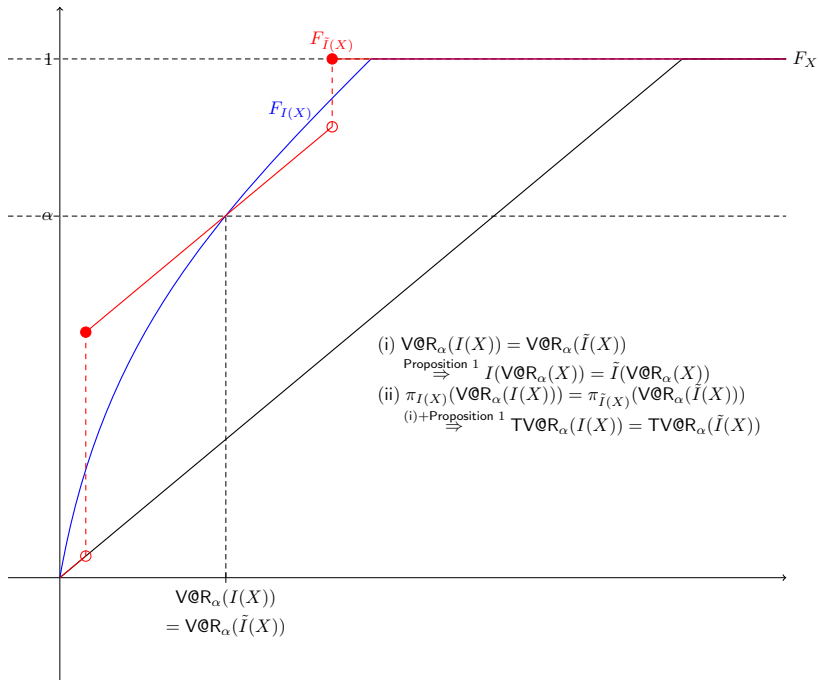


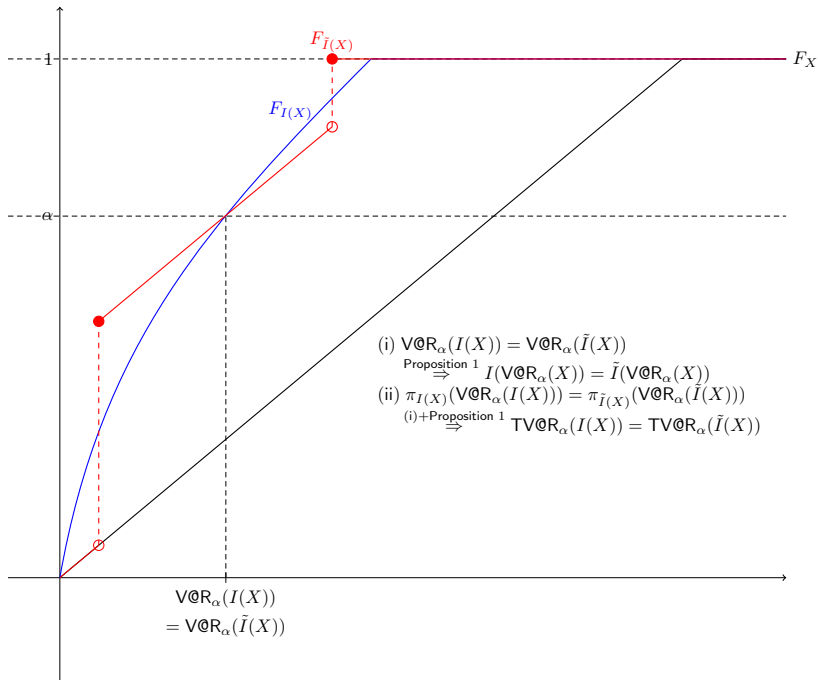


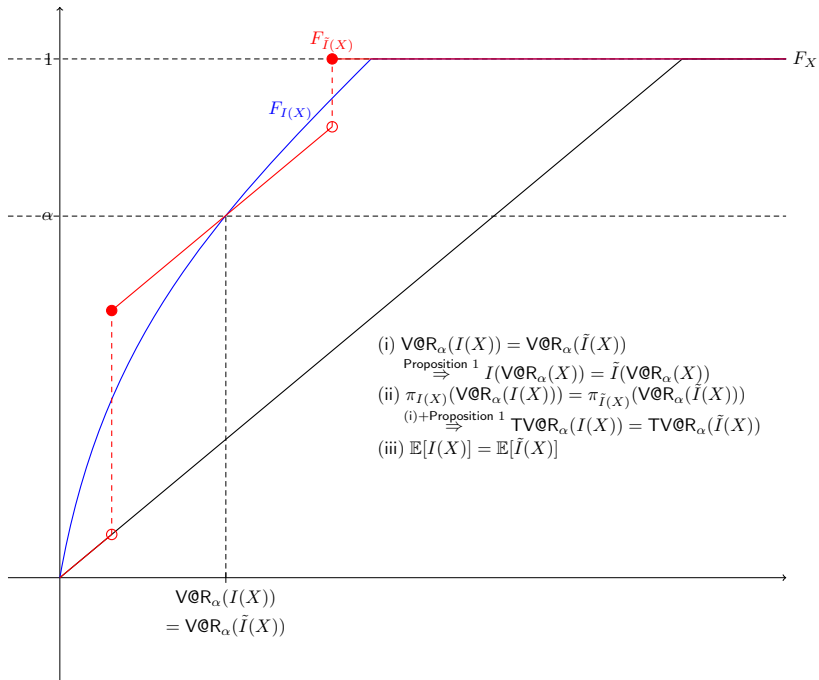


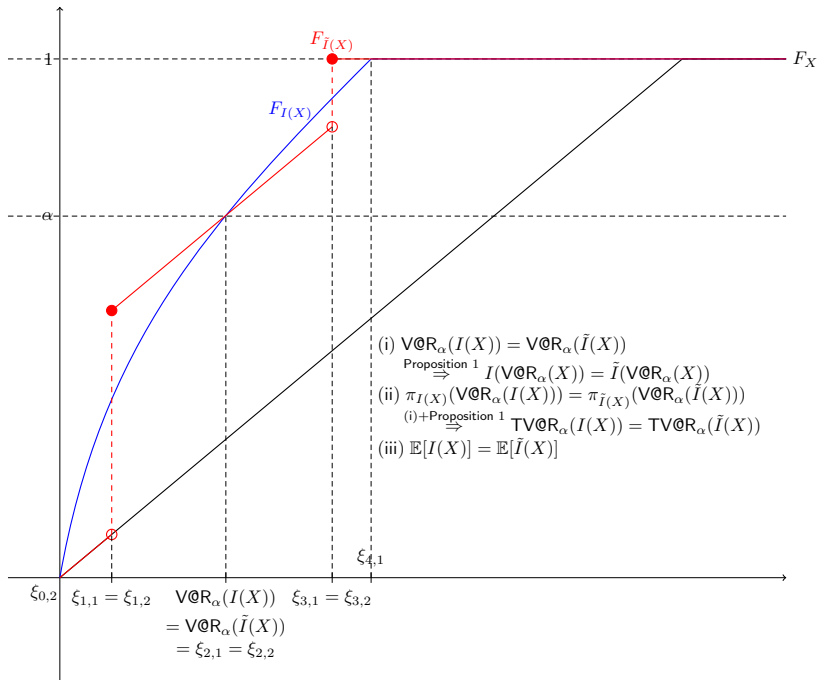


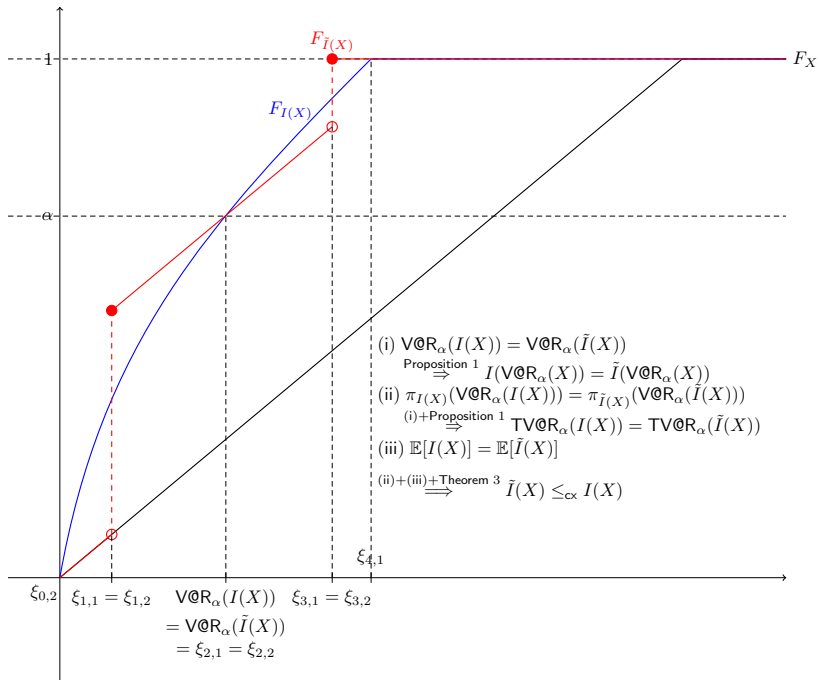


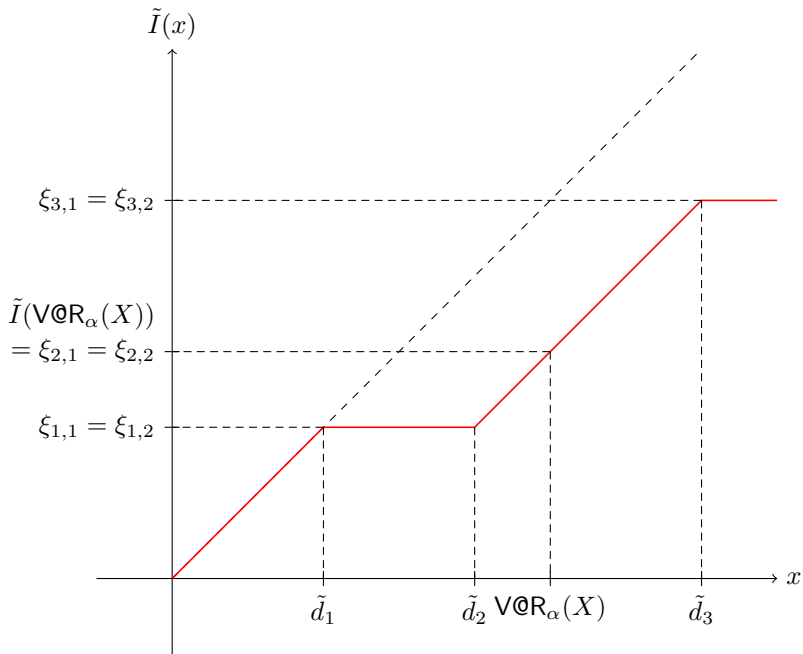




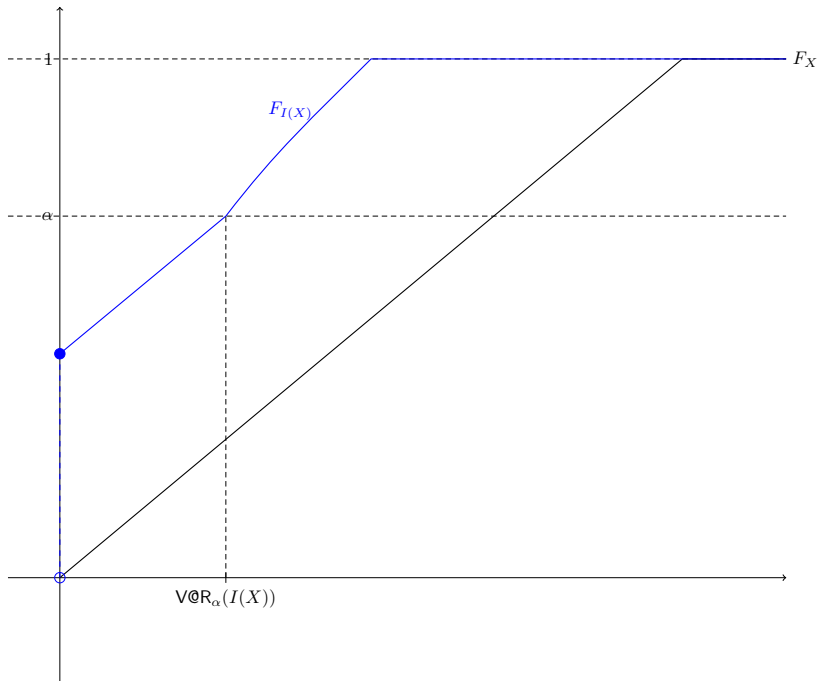


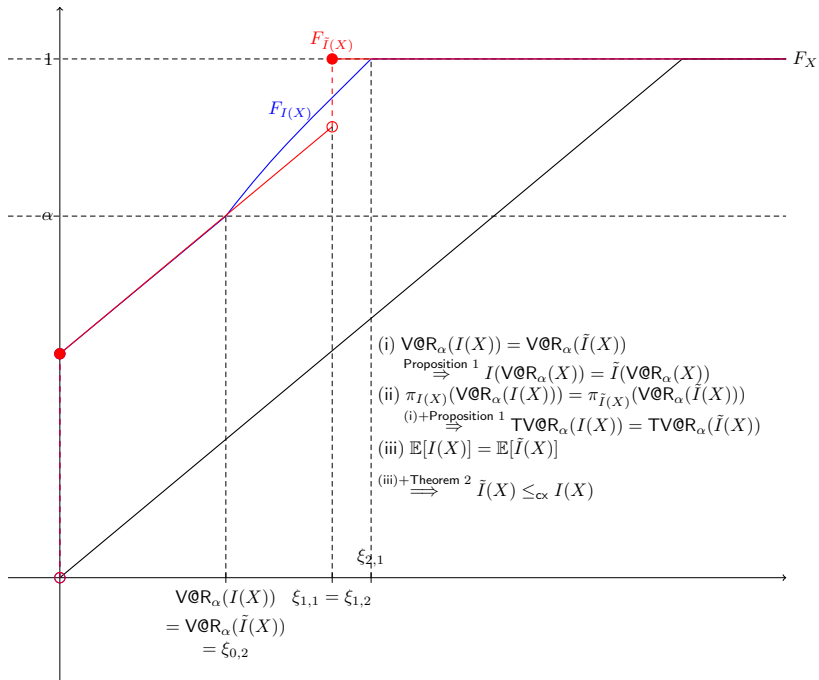


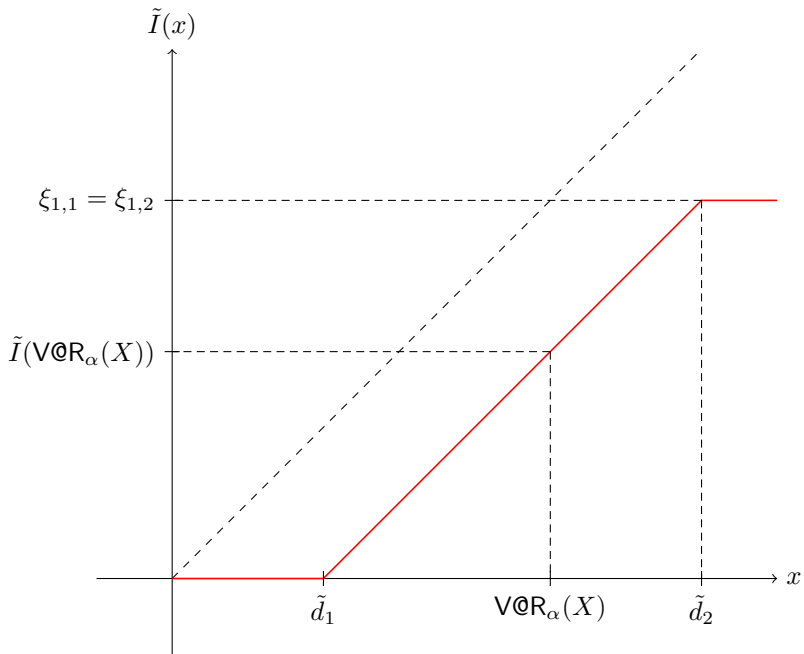


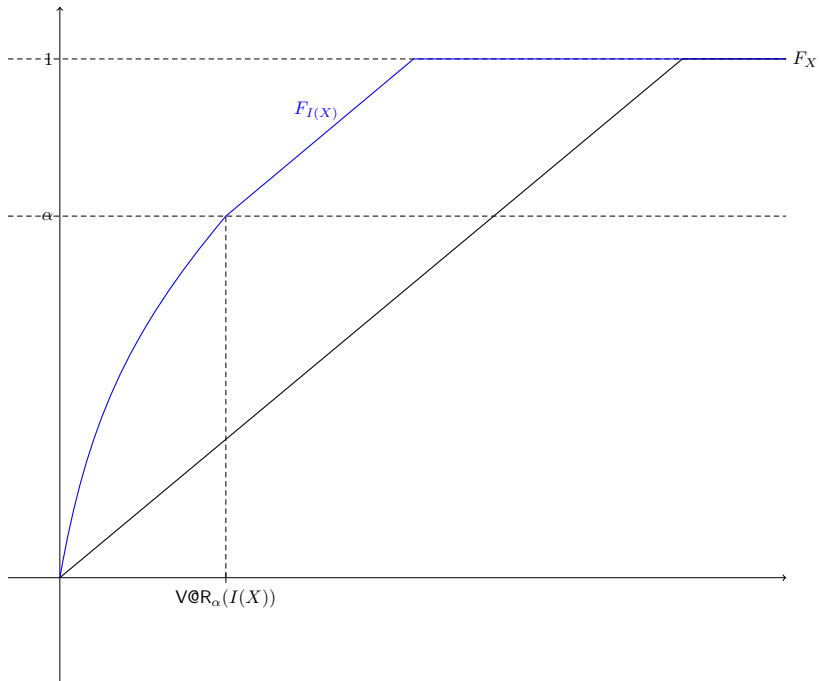


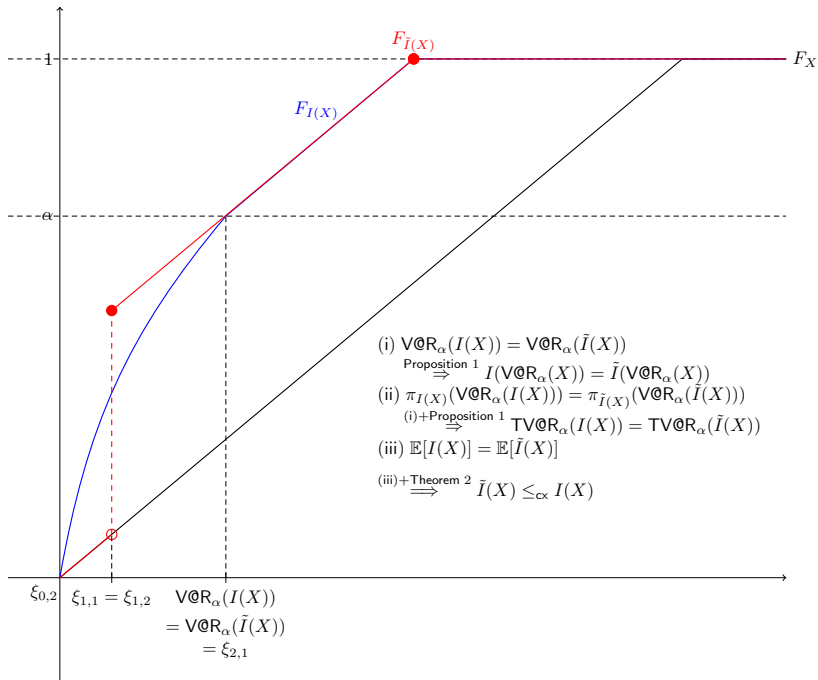


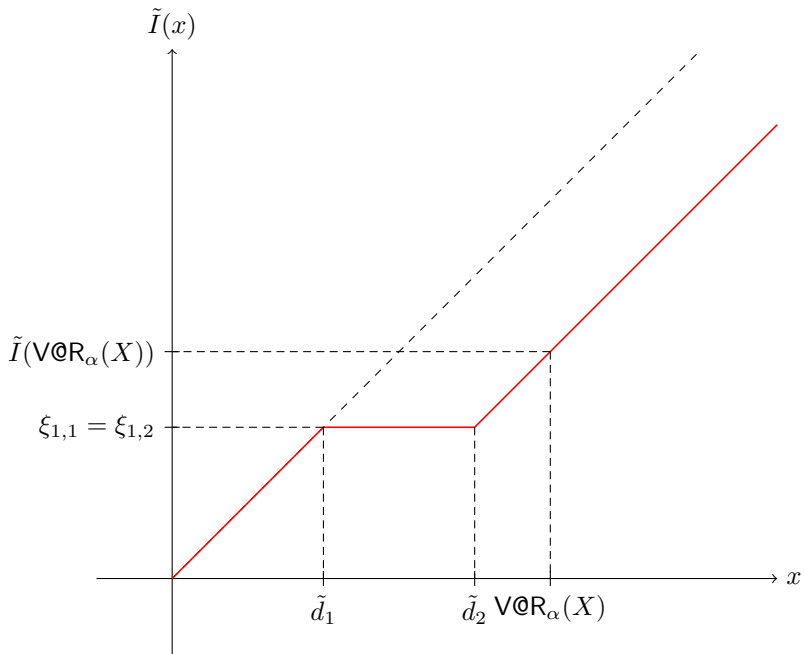












# Solution

To explicitly clarify the dependence of  $G$  on its three arguments, define the indicators:

$$J_i(G) \triangleq \begin{cases} 0 & \text{if } G \text{ does not depend on the } i\text{-th argument,} \\ 1 & \text{otherwise,} \end{cases}$$

for  $i = 1, 2, 3$ .

## Theorem 4

*Assume that  $G$  depends on all three arguments, such that  $J_1(G) = J_2(G) = J_3(G) = 1$ . The optimal ceded loss function for the problem takes the form of*

$$I^*(x) = x - (x - d_1^*)_+ + (x - d_2^*)_+ - (x - d_3^*)_+,$$

*for some  $0 \leq d_1^* \leq d_2^* \leq \text{V@R}_\alpha(X) \leq d_3^* \leq \text{ess sup } X$ .*

## Solution (cont.)

The construction of  $\tilde{I}$  such that

$$\text{TV@R}_\alpha(I(X)) = \text{TV@R}_\alpha(\tilde{I}(X))$$

is due to both

$$\text{V@R}_\alpha(I(X)) = \text{V@R}_\alpha(\tilde{I}(X)) \quad (1)$$

and, if  $\alpha < 1$ ,

$$\pi_{I(X)}(\text{V@R}_\alpha(I(X))) = \pi_{\tilde{I}(X)}(\text{V@R}_\alpha(\tilde{I}(X))). \quad (2)$$

Therefore, even if  $G$  does not depend on  $\text{V@R}$  but depends on  $\text{TV@R}$ , (1) is still needed to be held.

### Theorem 5 (Extension of Chi (2012))

*Assume that  $G$  depends only on the first and the third arguments, such that  $J_1(G) = J_3(G) = 1$  but  $J_2(G) = 0$ . The optimal ceded loss function for the problem takes the form of*

$$I^*(x) = x - (x - d_1^*)_+ + (x - d_2^*)_+ - (x - d_3^*)_+,$$

for some  $0 \leq d_1^* \leq d_2^* \leq \text{V@R}_\alpha(X) \leq d_3^* \leq \text{ess sup } X$ .



## Solution (cont.)

If  $G$  does not depend on  $\text{TV@R}$  but depends on  $\text{V@R}$ , (2) is not necessary anymore.

### Theorem 6 (Extension of Chi (2012))

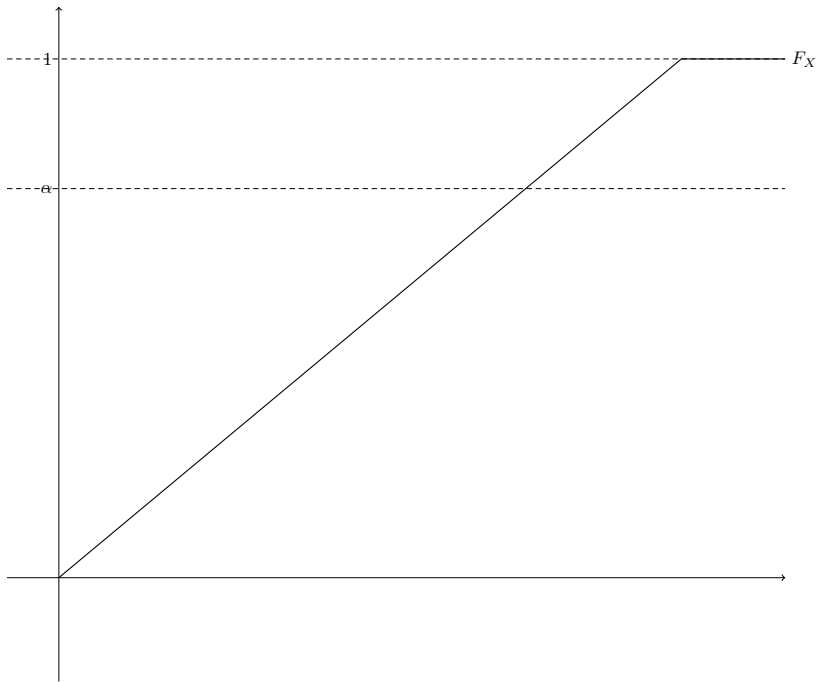
*Assume that  $G$  depends only on the first and the second arguments, such that  $J_1(G) = J_2(G) = 1$  but  $J_3(G) = 0$ . The optimal ceded loss function for the problem takes the form of either*

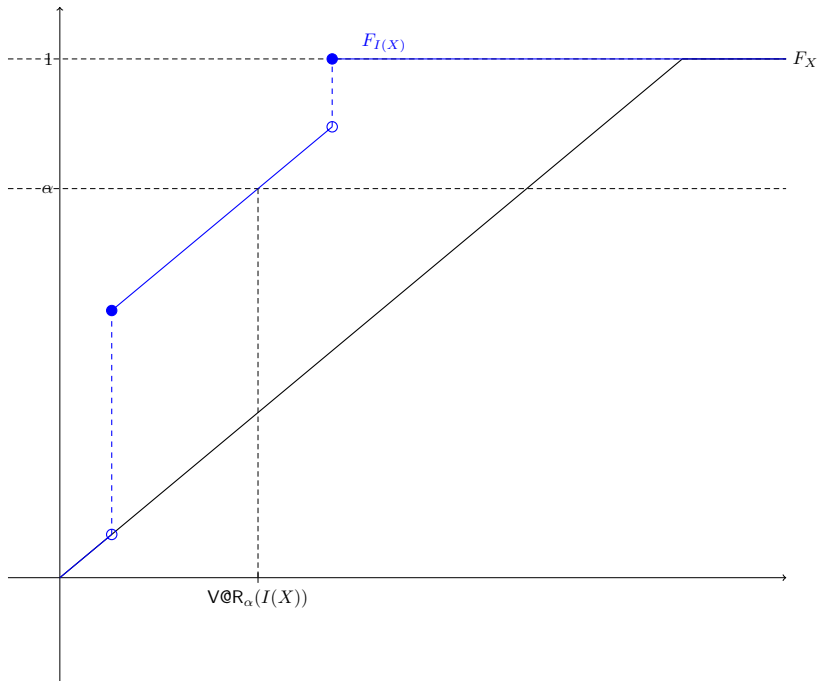
$$I^*(x) = x - (x - d_1^*)_+ + (x - d_2^*)_+ - (x - \text{V@R}_\alpha(X))_+,$$

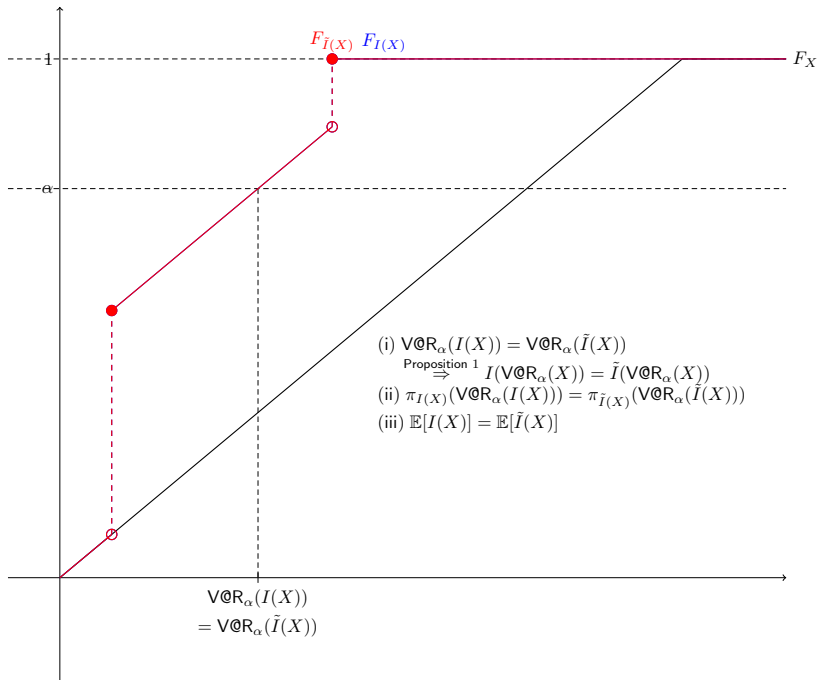
*for some  $0 \leq d_1^* \leq d_2^* \leq \text{V@R}_\alpha(X) \leq \text{ess sup } X$ , or*

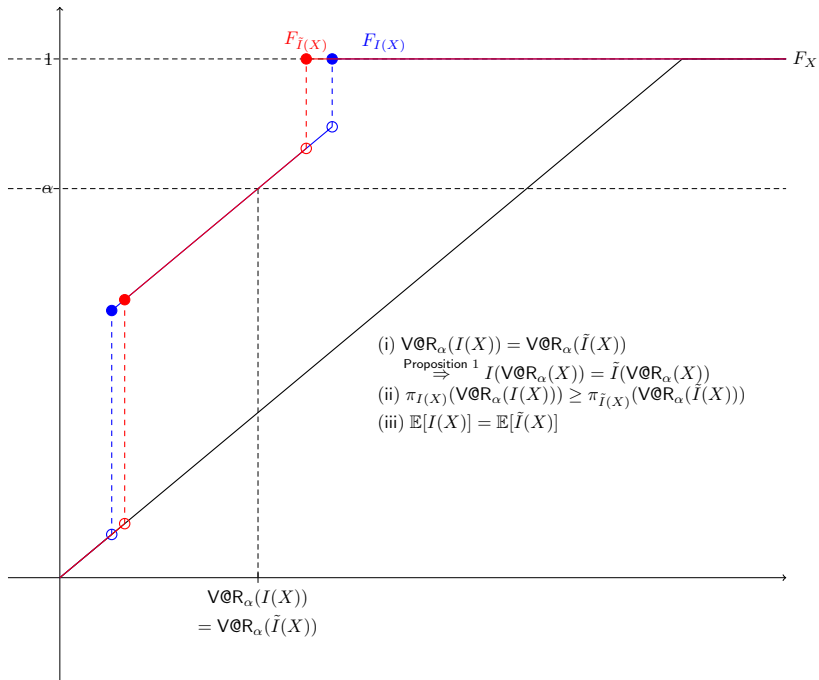
$$I^{**}(x) = x - (x - d^{**})_+,$$

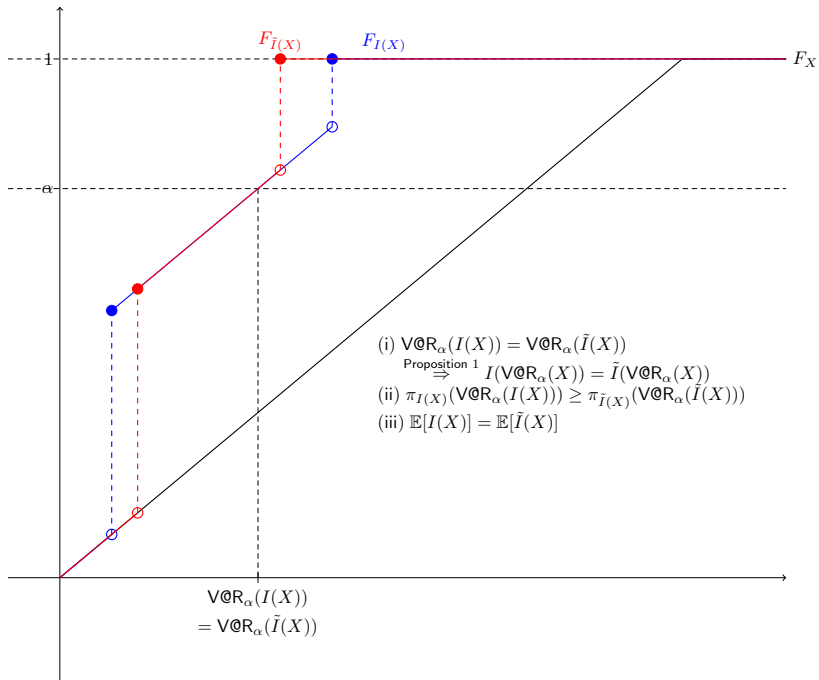
*for some  $0 \leq d^{**} \leq \text{ess sup } X$ .*

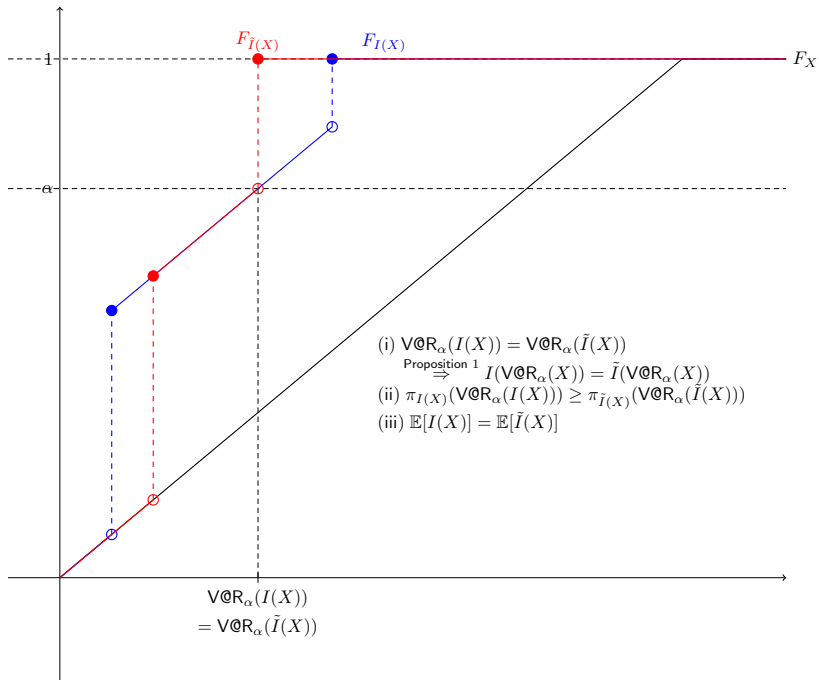


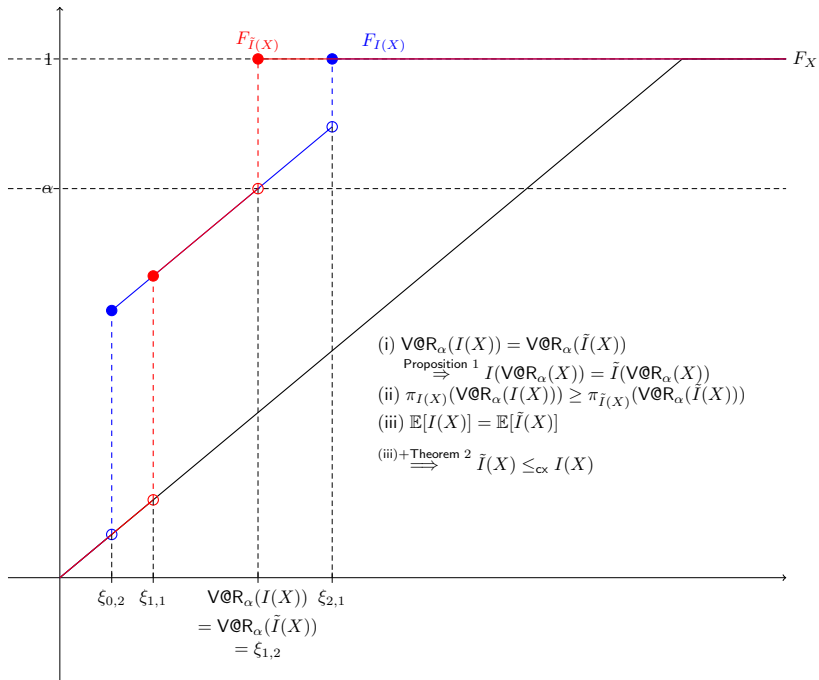




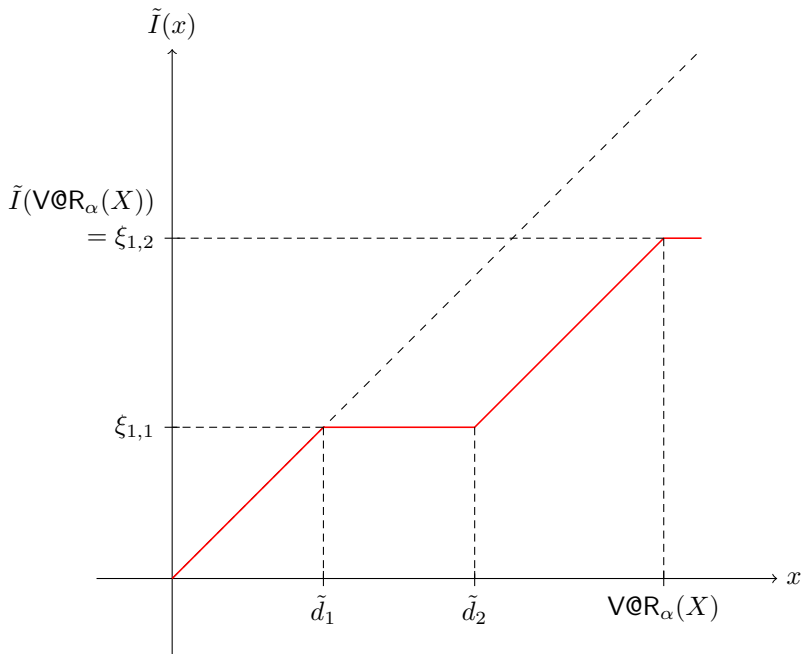


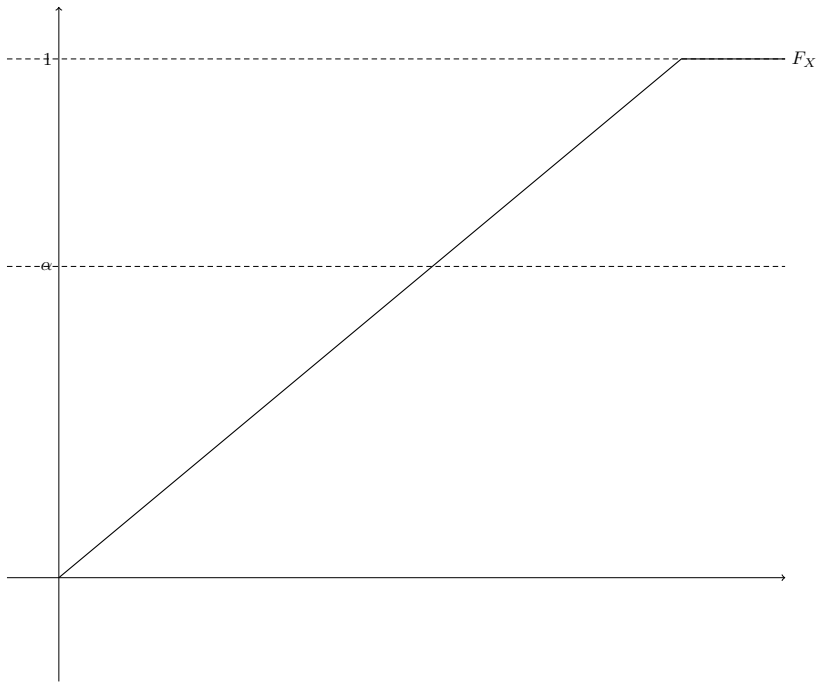


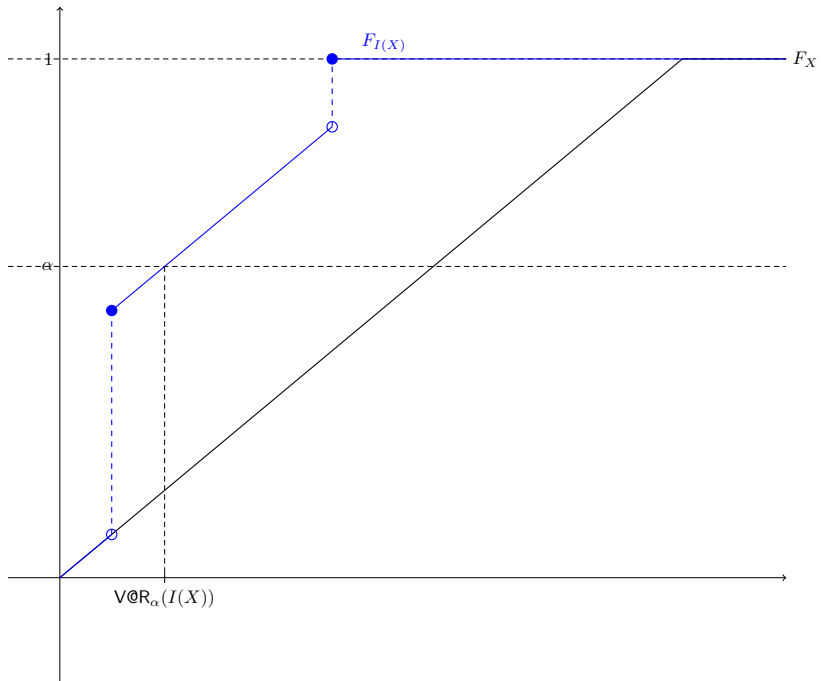


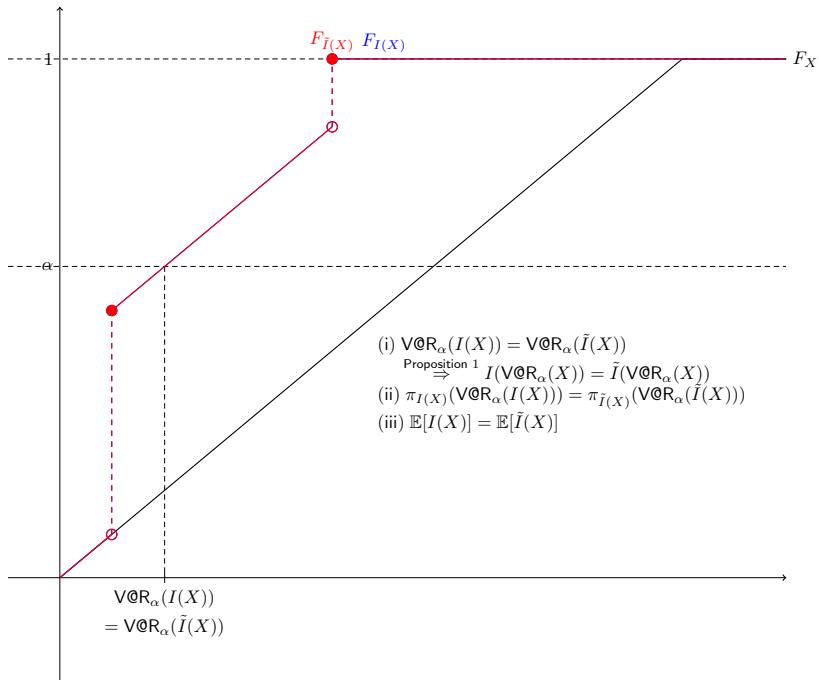


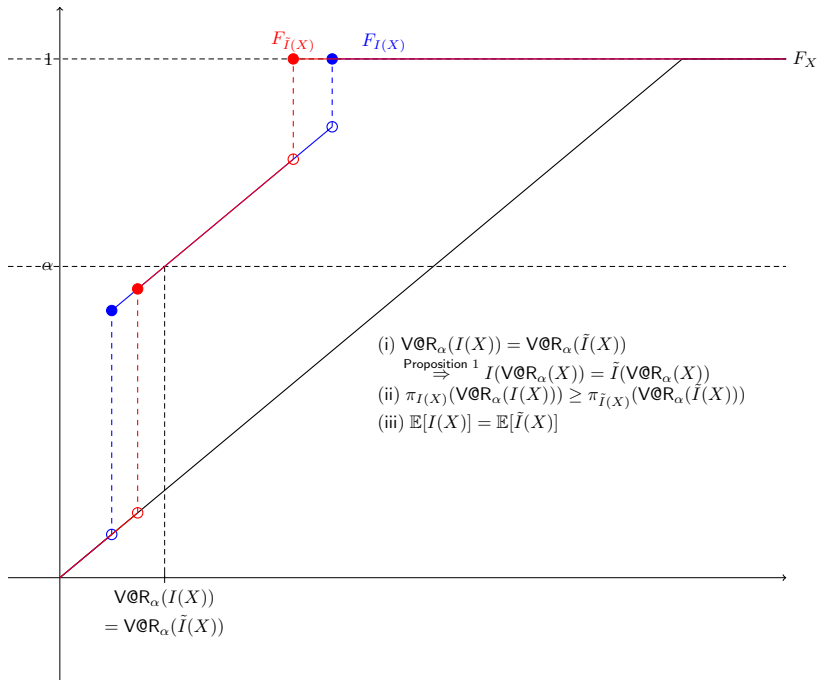


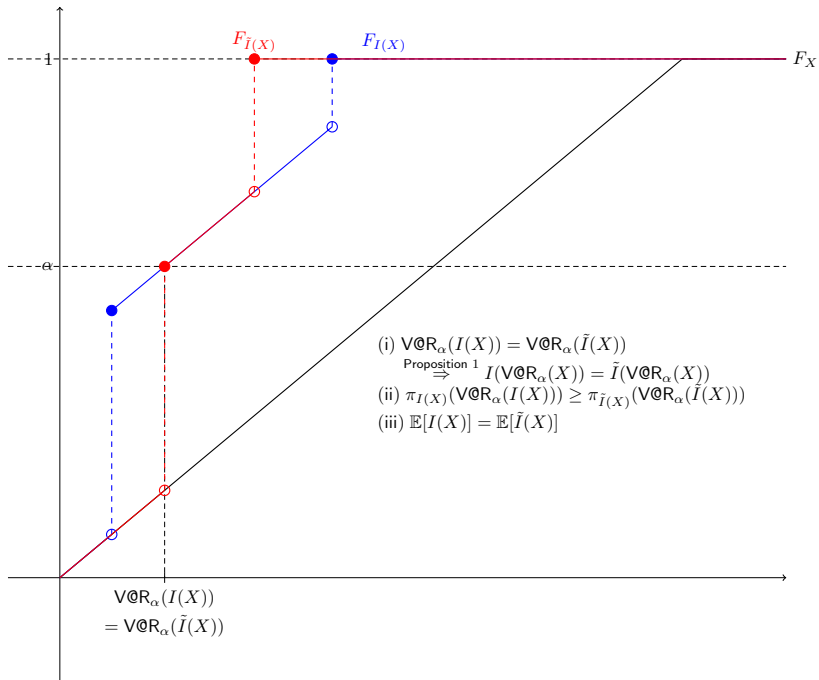


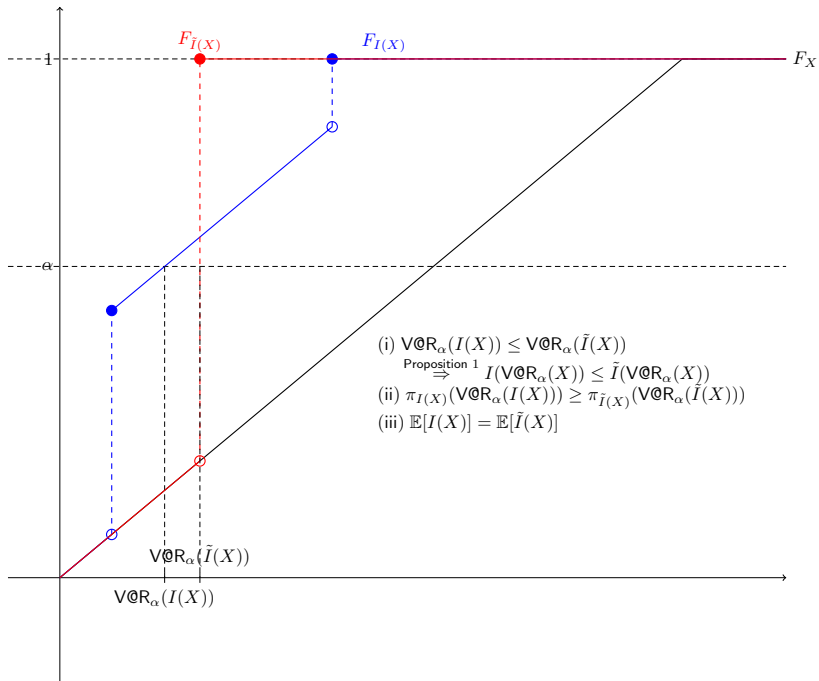


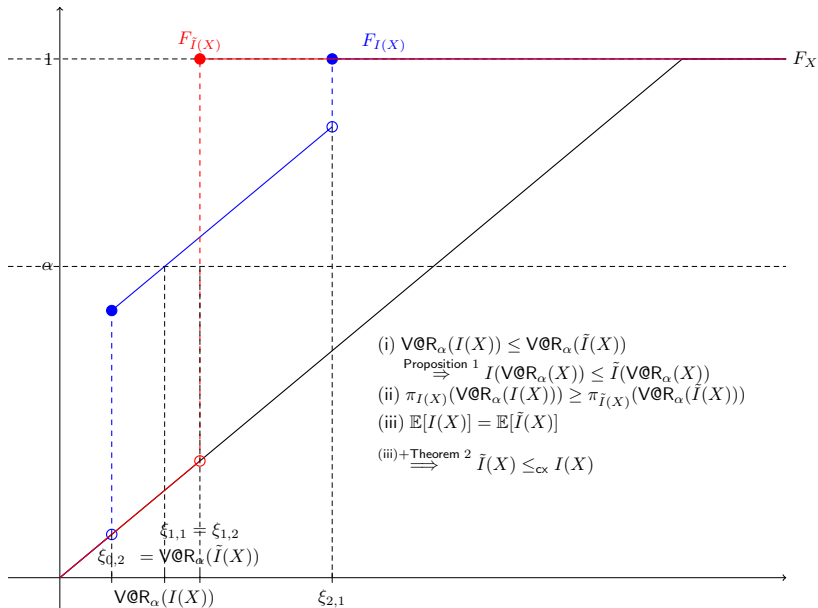




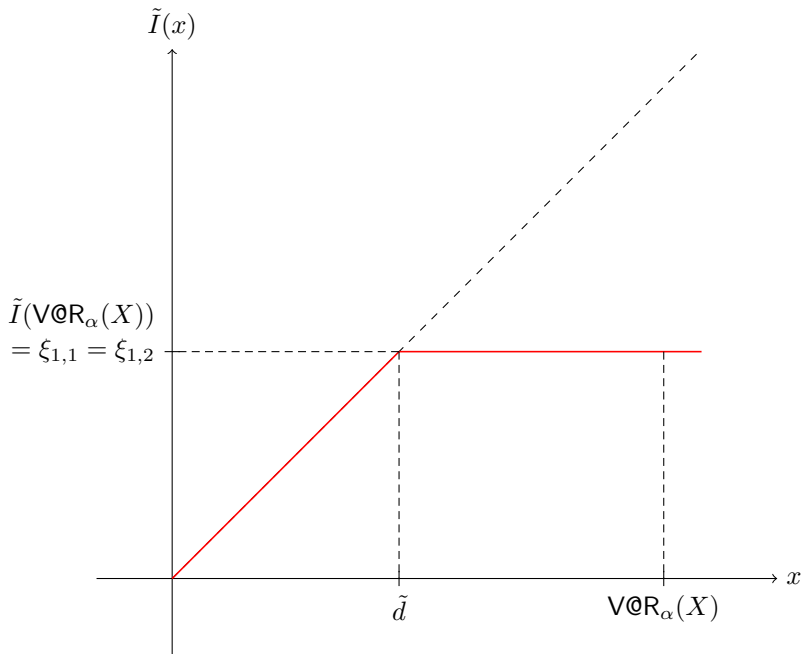


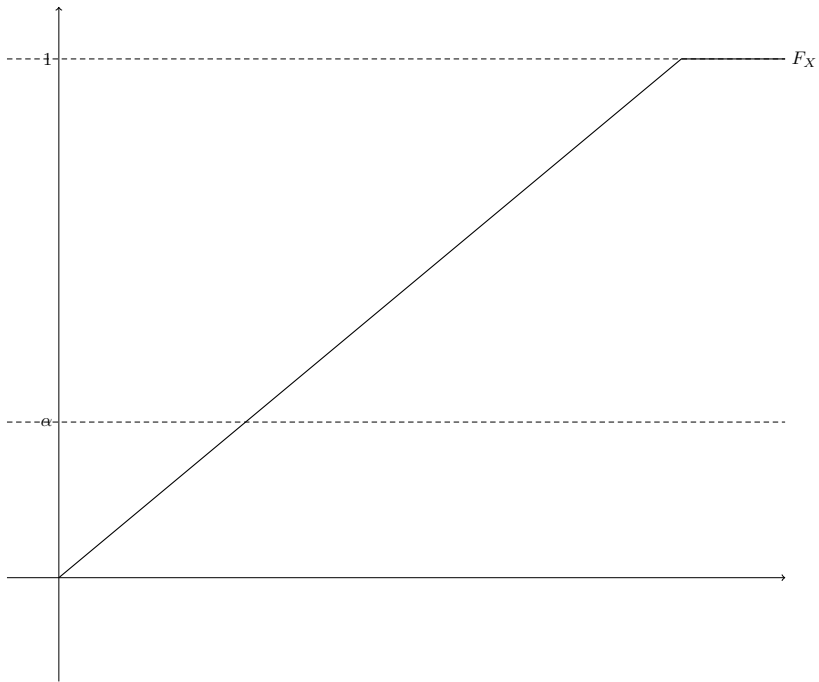


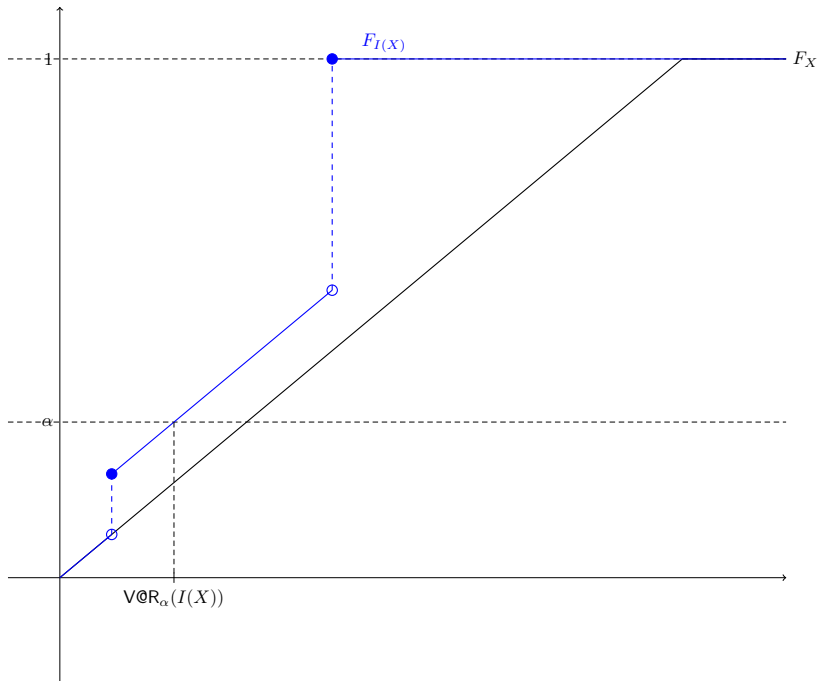


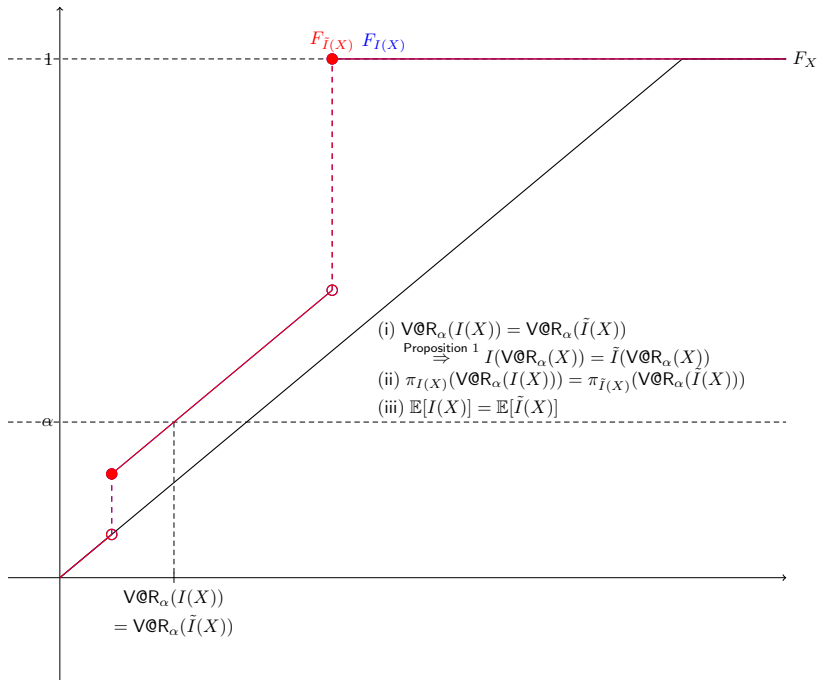


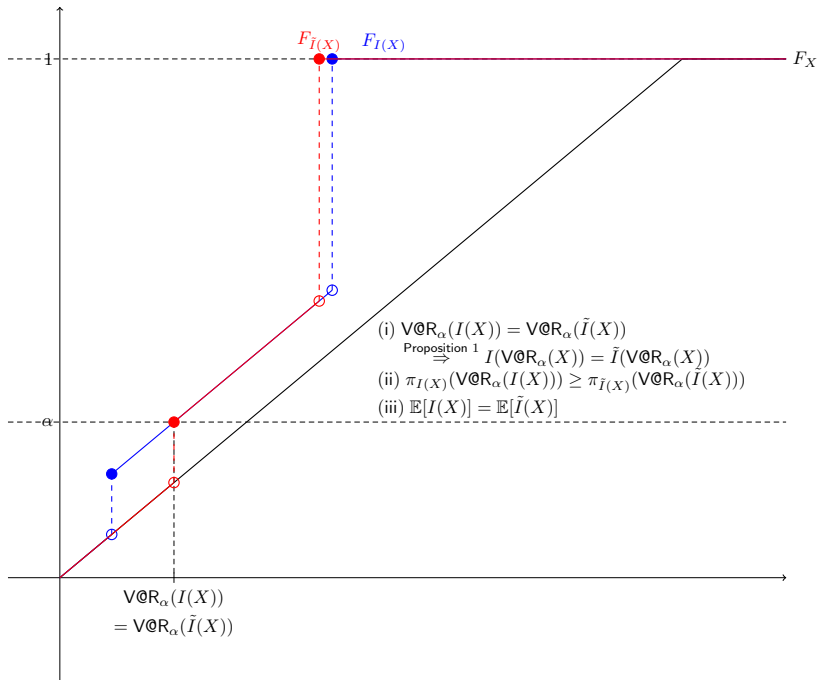


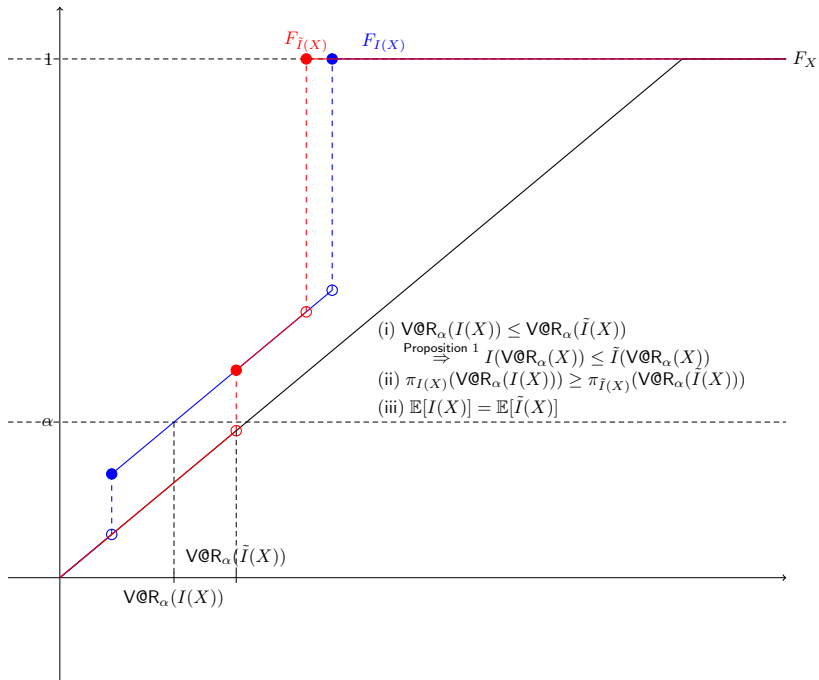


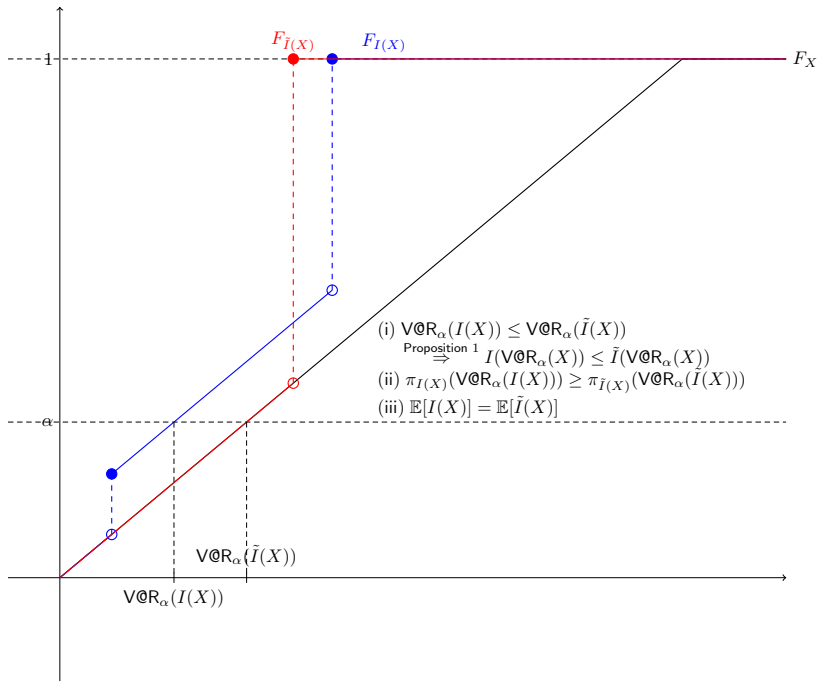


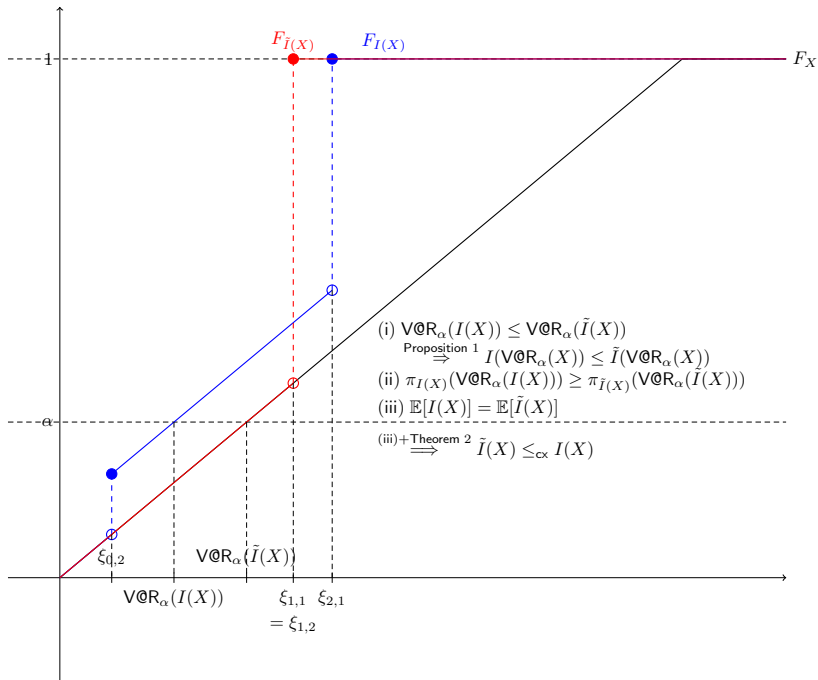




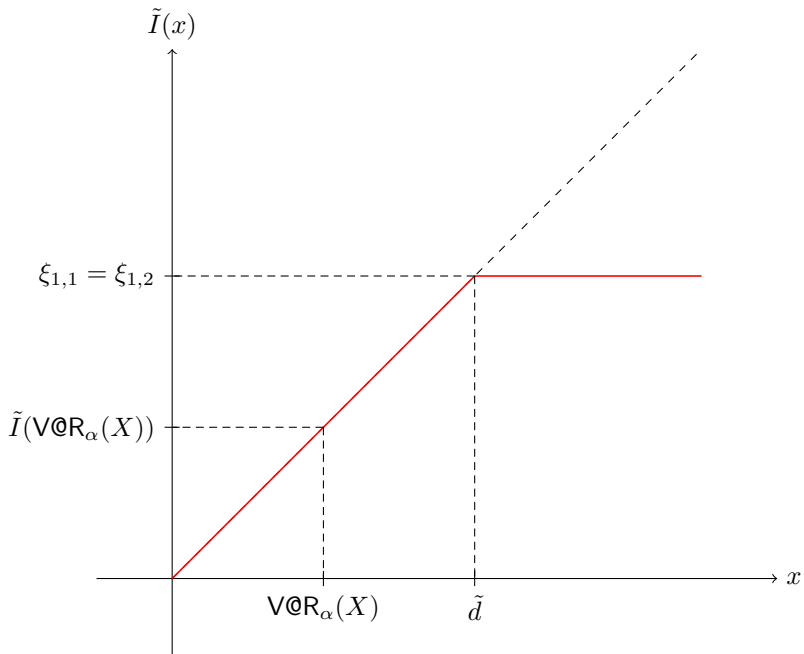










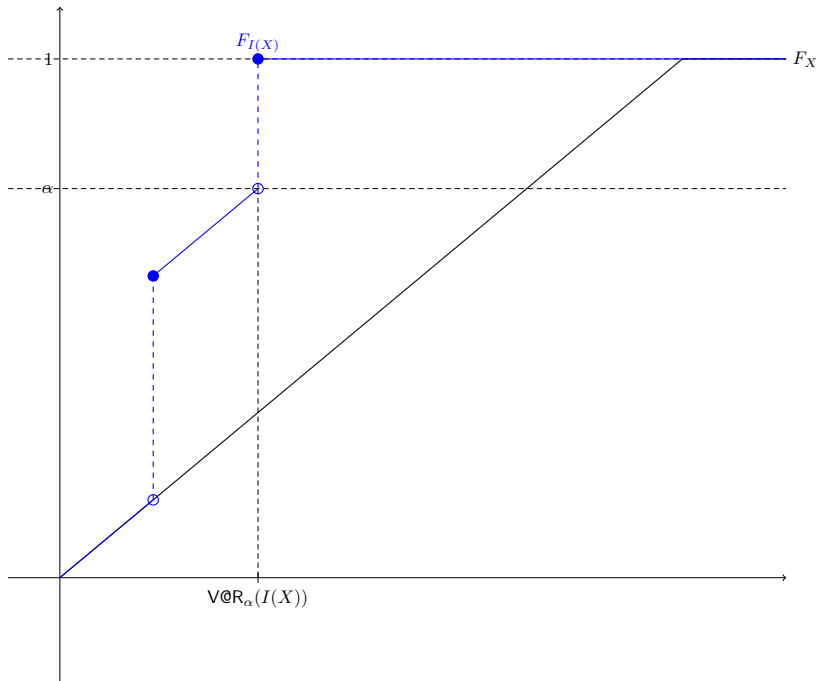


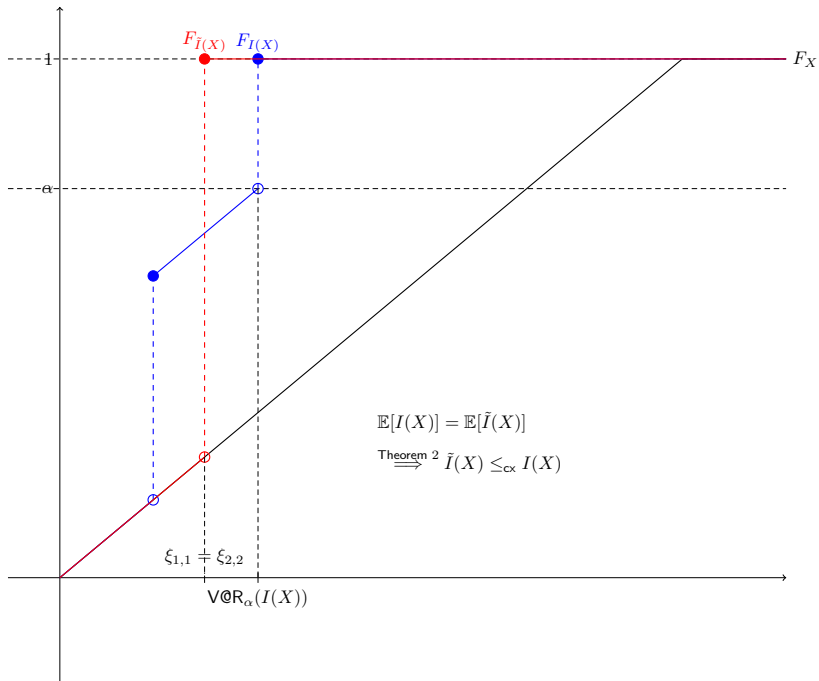
## Theorem 7

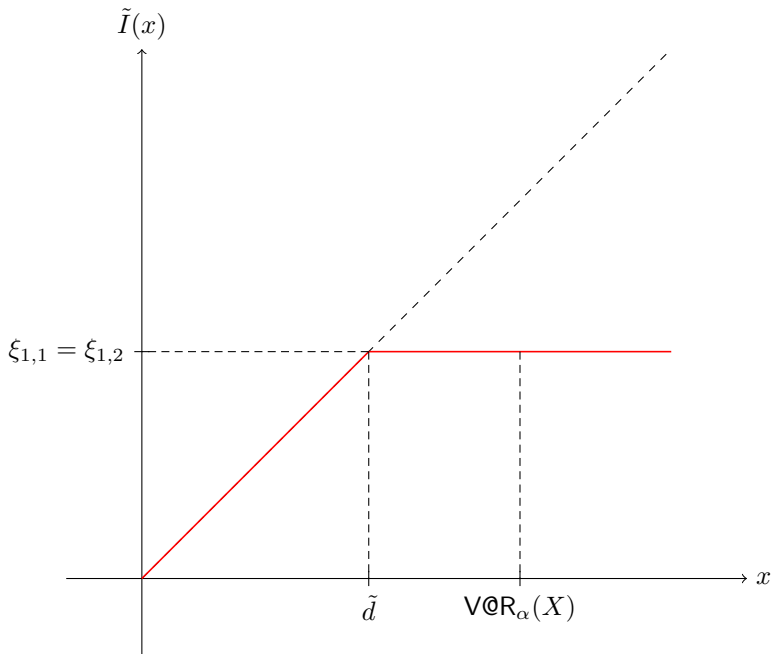
*Assume that  $G$  solely depends on the first argument, such that  $J_1(G) = 1$  but  $J_2(G) = J_3(G) = 0$ . The optimal ceded loss function for the problem takes the form of*

$$I^*(x) = x - (x - d^*)_+,$$

*for some  $0 \leq d^* \leq \text{ess sup } X$ .*



















# Conclusion and Future Potential Works

- We studied a broad class of optimal insurance decision problems, in which the objective involves the expectation,  $V@R$  and  $TV@R$  of the total loss exposure of the policyholder, while the premium principle of the insurance indemnity is characterized by its expectation and a functional preserving convex order, which includes well-known premium principles, such as the Wang's principle.
- We applied the Karlin-Novikoff-Stoyan-Taylor multiple-crossing conditions for convex ordering to solve the most general problem.
- We showed that the optimal solution of these problems in general takes the double layer form.
- We further solved sub-problems by making use of the solution of the most general problem and the Karlin-Novikoff once-crossing conditions for convex ordering.
- Future potential works: general risk measure in objective,  $V@R$  and  $TV@R$  of the insurer in premium, functional preserving other stochastic ordering in premium, ...

# References






-  Chi, Y. (2012): Reinsurance arrangements minimizing the risk-adjusted value of an insurers liability, *ASTIN Bulletin* 42(2), 529–557.
-  Denuit, M., Dhaene, J., Goovaerts, M., and Kaas, R. (2005): *Actuarial Theory for Dependent Risks: Measures, Orders and Models*. John Wiley & Sons.
-  Dhaene, J., Denuit, M., Goovaerts, M. J., Kaas, R., and Vyncke, D. (2002): The concept of comonotonicity in actuarial science and finance: theory, *Insurance: Mathematics and Economics* 31(1), 3–33.
-  Dhaene, J., Vanduffel, S., Goovaerts, M.J., Kaas, R., Tang, Q., and Vyncke, D. (2006): Risk measures and comonotonicity: a review, *Stochastic Models* 22(4), 573–606.
-  Gollier, C. and Schlesinger, H. (1996): Arrows theorem on the optimality of deductibles: a stochastic dominance approach, *Economic Theory* 7(2), 359–363.
-  Hürlimann, W. (1998): On stop-loss order and the distortion pricing principle, *ASTIN Bulletin* 28, 119–134.

# References

-  Hürlimann, W. (2008a): Extremal moment methods and stochastic orders: Application in actuarial science., *Boletín de la Asociación Matemática Venezolana* 15(1), 5–110. Chapters I, II and III.
-  Hürlimann, W. (2008a): Extremal moment methods and stochastic orders: Application in actuarial science., *Boletín de la Asociación Matemática Venezolana* 15(2), 153–301. Chapters IV, V and VI.
-  Kaas, R., Goovaerts, M., Dhaene, J., and Denuit, M. (2008): *Modern Actuarial Risk Theory: Using R*. Springer Science & Business Media.
-  Kaas, R., van Heerwaarden, A. E., and Goovaerts, M. J. (1994): *Ordering of Actuarial Risks*. CAIRE. Brussels.
-  Karlin, S. and Novikoff, A. (1963): Generalized convex inequalities, *Pacific Journal of Mathematics* 13(4), 1251–1279.
-  Müller, A. and Stoyan, D. (2002): *Comparison Methods for Stochastic Models and Risks*. John Wiley & Sons.



# References

-  Ohlin, J. (1969): On a class of measures of dispersion with application to optimal reinsurance, *ASTIN Bulletin* 5(2), 249–266.
-  Rüschemdorf, L. (2013): *Mathematical Risk Analysis: Dependence, Risk Bounds, Optimal Allocations and Portfolios*. Springer-Verlag Berlin Heidelberg.
-  Shaked, M. and Shanthikumar, J. G. (2007): *Stochastic Orders*. Springer Science & Business Media.
-  Stoyan, D. (1983): *Comparison Methods for Queues and Other Stochastic Models*. John Wiley & Sons.
-  Taylor, J. M. (1983): Comparisons of certain distribution functions, *Statistics: A Journal of Theoretical and Applied Statistics* 14(3), 397–408.

**Thank you!**